CANDIDATE CLUSTERS OF GALAXIES AT Z > 1.3 IDENTIFIED IN THE SPITZER SPT DEEP FIELD SURVEY
A. Rettura$^{1,2}$, J. Martinez-Manso$^3$, D. Stern$^4$, S. Mei$^{4,5,6}$, M.L.N. Ashby$^7$, M. Brodwin$^8$, D. Gettings$^9$, A.H. Gonzalez$^2$, S.A. Stanford$^{10,11}$ and J.G. Bartlett$^{11,1}$

ABSTRACT
We present 279 galaxy cluster candidates at $z > 1.3$ selected from the $94\,\text{deg}^2$ Spitzer South Pole Telescope Deep Field (SSDF) survey. We use a simple algorithm to select candidate high-redshift clusters of galaxies based on Spitzer/IRAC mid-infrared data combined with shallow all-sky optical data. We identify distant cluster candidates in SSDF adopting an overdensity threshold that results in a high purity (80%) cluster sample based on tests in the Spitzer Deep, Wide-Field Survey of the Boötes field. Our simple algorithm detects all three $1.4 < z \leq 1.75$ X-ray detected clusters in the Boötes field. The uniqueness of the SSDF survey resides not just in its area, one of the largest contiguous extragalactic fields observed with Spitzer, but also in its deep, multi-wavelength coverage by the South Pole Telescope (SPT), Herschel/SPIRE and XMM-Newton. This rich dataset will allow direct or stacked measurements of Sunyaev-Zel’dovich effect decrements or X-ray masses for many of the SSDF clusters presented here, and enable systematic study of the most distant clusters on an unprecedented scale. We measure the angular correlation function of our sample and find that these candidates show strong clustering. Employing the COSMOS/UltraVista photometric catalog in order to infer the redshift distribution of our cluster selection, we find that these clusters have a comoving number density $n_c = (0.7^{+0.3}_{-0.4}) \times 10^{-7} h^3 \text{Mpc}^{-3}$ and a spatial clustering correlation scale length $r_0 = (32 \pm 7) h^{-1} \text{Mpc}$. Assuming our sample is comprised of dark matter halos above a characteristic minimum mass, $M_{\text{min}}$, we derive that at $z = 1.5$ these clusters reside in halos larger than $M_{\text{min}} = 1.5^{+0.9}_{-0.7} \times 10^{14} h^{-1} M_\odot$. We find the mean mass of our cluster sample to be equal to $M_{\text{mean}} = 1.9^{+1.0}_{-0.8} \times 10^{14} h^{-1} M_\odot$, thus our sample contains the progenitors of present-day massive galaxy clusters.

Subject headings: galaxies: clusters: general — galaxies: high-redshift — galaxies: statistics — cosmology: observations — cosmology: large-scale structure of universe — infrared: galaxies

1. INTRODUCTION

Emerging from the cosmic web, galaxy clusters are the most massive gravitationally bound structures in the universe. Thought to have begun their assembly at $z > 2$, clusters provide insights into the growth of large-scale structure as well as the physics that drives galaxy evolution. Understanding how and when the most massive galaxies assemble their stellar mass, stop forming stars, and acquire their observed morphologies remain outstanding questions. The redshift range $1.4 < z < 2$ is a key epoch in this respect: elliptical galaxies start to become the dominant population in cluster cores, and star formation in spiral galaxies is being quenched (e.g., Blakeslee et al. 2006; Rosati et al. 2009; Mei et al. 2009; Overzier et al. 2010; 2011; Raichoor et al. 2011; Strazzullo et al. 2010; Stanford et al. 2012; Snyder et al. 2012; Zeimann et al. 2012; Mei et al. 2012; Nantais et al. 2013). Interestingly, some field galaxy studies find that the star formation rate (SFR)-density relation reverses at $z = 1$ relative to $z = 0$ (Cooper et al. 2007; Eilbaz et al. 2007), such that star formation no longer decreases with increasing galaxy density at $z = 1$. However some other studies disagree with this result (Patel et al. 2009; Muzzin et al. 2012; Scoville et al. 2013) and conclude the reversal must happen at $z > 1$ as they find the local density correlations to be already in place by $z = 1$. There is also observational evidence for a progressive increase in the amount of star formation that occurs in galaxy cluster cores at $z \gtrsim 1.4$ (e.g., Hilton et al. 2009; Hayashi et al. 2010; Tran et al. 2010; Passhender et al. 2011; Tadaki et al. 2012; Brodwin et al. 2013; Alberts et al. 2014). This suggests that significant star formation is occurring in high-density environments at early epochs. Therefore, increasing evidence points to clusters at $1.5 < z < 2$ as
being the ideal laboratories to study cluster formation and to catch in the act transformations in their stellar populations.

Until recently, however, this redshift range was essentially unreachable with available instrumentation, with clusters at these redshifts exceedingly challenging to identify from either ground-based optical/near-infrared (NIR) imaging or from X-ray surveys. Mid-infrared (MIR) imaging with Spitzer has changed the landscape. Previous Spitzer wide-area surveys have proven effective at identifying samples of galaxy clusters down to low masses at $1 < z < 2$ (e.g., SDWFS, SWIRE, CARLA [Eisenhardt et al. 2003], Papovich [2008], Wilson et al. 2009, Demarco et al. 2010, Galametz et al. 2010, Stanford et al. 2012, Zeimann et al. 2012, Brodwin et al. 2013, Galametz et al. 2013, Muzzin et al. 2013, Wylezalek et al. 2013) where current X-ray observations are restricted to only the most massive systems. X-ray follow-up has verified several of these MIR-selected clusters, implying masses of a few $10^{14} M_\odot$ (Papovich et al. 2010, Brodwin et al. 2011, Muzzin et al. 2012; Wylezalek et al. 2013) where current X-ray surveys are restricted to only the most massive systems. X-ray follow-up has verified several of these MIR-selected clusters, implying masses of a few $10^{14} M_\odot$ (Papovich et al. 2010, Brodwin et al. 2011, Muzzin et al. 2012; Wylezalek et al. 2013) where current X-ray observations are restricted to only the most massive systems. X-ray follow-up has verified several of these MIR-selected clusters, implying masses of a few $10^{14} M_\odot$ (Papovich et al. 2010, Brodwin et al. 2011, Muzzin et al. 2012; Wylezalek et al. 2013) where current X-ray observations are restricted to only the most massive systems. X-ray follow-up has verified several of these MIR-selected clusters, implying masses of a few $10^{14} M_\odot$ (Papovich et al. 2010, Brodwin et al. 2011, Muzzin et al. 2012; Wylezalek et al. 2013) where current X-ray observations are restricted to only the most massive systems.

We select candidate distant galaxy clusters based on the 4.5 $\mu$m-selected Spitzer/IRAC band-merged ([3.6], [4.5]) catalog, which contains $\sim 3.7$ million distinct sources down to the SSDF $5\sigma$ sensitivity limit of 21.46 AB mag (9.4 $\mu$Jy) at 4.5 $\mu$m; the corresponding 5$\sigma$ sensitivity of the 3.6 $\mu$m bandpass is 21.79 AB mag (7.0 $\mu$Jy). Throughout this paper we use aperture-corrected, $4''$-diameter aperture magnitudes for the IRAC data.

2. SuperCOSMOS Optical Data

The SuperCOSMOS survey (Hambly et al. 2001) provides all-sky optical photometry based on scans of photographic Schmidt survey plates from the UK Schmidt Telescope (UKST) Southern Surveys (Hartley & Dawe 1983, Cannon 1984) and Palomar Oschin Schmidt Telescope Surveys (POSS: Minkowski & Abell 1963, Reid et al. 1991). These shallow data provide $I$-band magnitudes down to $I \sim 20.45$ mag (AB) in the SSDF (Hambly et al. 2001). Portions of the SSDF field have deeper optical data from more recent surveys, such as the Blanco Cosmology Survey which reaches $\sim 1\mu$Jy depths in $griz$ (Desai et al. 2012), and the whole field will be covered by the Dark Energy Survey (Mohr et al. 2012). However, in the interest of uniformity in our cluster search over the widest possible area, we only consider the relatively shallow SuperCOSMOS optical data in the following analysis.

3. IDENTIFYING HIGH-REDSHIFT GALAXY CLUSTERS

3.1. Methodology

We select candidate distant galaxy clusters based on their [3.6] − [4.5] galaxy color, following the approach of Papovich (2008); that methodology was proven effective by discovering a $z = 1.62$ galaxy cluster in the SWIRE-XMM field (Lonsdale et al. 2003) using data of similar depth to what is available in the SSDF (Papovich et al. 2010). The method takes advantage of the fact that [3.6] − [4.5] color is a linear function of redshift between $0.7 \lesssim z \lesssim 1.5$ (Fig. 1), and thus can be used as an effective redshift indicator. At $z \gtrsim 1.5$ the color reaches a plateau out to $z \sim 3$. The basis of the selection is that galaxy stellar populations with ages $> 10$ Myr have a prominent bump at $\sim 1.6$ $\mu$m due to a minimum in the opacity of the H− ion present in the atmospheres of cool stars. This feature is seen in the spectral energy distribution (SED) of essentially all galaxies, largely independent of star formation history or age (Jolfl 1988, Simpson & Eisenhardt 1999, Sawick 2002, Sorba & Sawicki 2010). However, above redshift $z \sim 1.5$ the color is not precise enough to be useful to estimate redshifts other than to constrain the redshift to be larger than $z \gtrsim 1.5$.
High Redshift Clusters in the Spitzer SPT Deep Field Survey

The evolution with redshift of SuperCOSMOS $I$-band magnitude (top panel), $[3.6] - [4.5]$ IRAC color (middle panel) and $[4.5]$ IRAC magnitude (bottom panel) for Bruzual & Charlot (2003) simple stellar population models formed at formation redshifts, $z_f$, of 5 (black dot-dashed line). The IRAC color selection criterion we adopt is optimized to find galaxies with $z > 1.3$. The $[3.6] - [4.5]$ color serves as a good redshift indicator for $0.7 < z < 1.5$; above this redshift, the $[3.6] - [4.5]$ color evolution with redshift flattens. To alleviate a known contamination from foreground interlopers at $z \sim 0.3$ (e.g., Papovich 2008; Muzzin et al. 2013) we also apply two magnitude cuts ($[4.5] > 19.5$ and $I > 20.45$).

As shown in the middle panel of Fig. 1, while an IRAC color cut $[3.6] - [4.5] > -0.1$ is effective at distinguishing galaxies at $z > 1.3$, having at least one relatively shallow optical band is very useful for alleviating contamination from foreground interlopers at $z \sim 0.3$ (see discussion in Muzzin et al. 2013). To this aim, we apply magnitude cuts in the $I$ and $[4.5]$ bands to remove most $z < 0.4$ galaxies (see bottom and top panel of Fig. 1). Combined, these cuts effectively remove the bulk of the foreground galaxy population at $z < 1.3$. The remaining sources consist predominantly of high-redshift galaxies. We note that other sources of contamination are cool brown dwarfs (Stern et al. 2007) and powerful AGNs at all redshifts (Stern et al. 2005); however, these are expected not to be contaminants (see discussion in Galametz et al. 2012).

Therefore, to discover distant clusters of galaxies at $z > 1.3$ we have implemented a simple three-filter algorithm to search for overdensities of galaxies based upon their Spitzer IRAC color ($[3.6] - [4.5] > -0.1$), their 4.5$\mu$m magnitude ($[4.5] > 19.5$) and requiring non-detection in the SuperCOSMOS $I$-band data ($I > 20.45$). Similar algorithms have been demonstrated to be effective by several programs (Papovich et al. 2010; Galametz et al. 2010, 2013; Gettings et al. 2012; Muzzin et al. 2013).

Fig. 1.— Distribution of the completeness-corrected excess number of objects with respect to the local background, $\Delta N$, with $[3.6] - [4.5] > -0.1$, $19.5 < [4.5] < 21.46$ and $I > 20.45$, within 1.0′ radius from each individual source in the SSDF catalog fulfilling the aforementioned color criteria. The mean number of excess objects is $\langle \Delta N \rangle = 0.4 \pm 3.7$. The red curve shows a Gaussian fit (iteratively clipping at 2σ). The dashed line indicates the minimum detection significance threshold adopted here of 5.2 times the standard deviations of the Gaussian distribution, corresponding to objects with more than 19.5 similarly selected (completeness corrected) companions within 1.0′.

Fig. 2.— Spatial distribution of 279 galaxy cluster candidates with $X_f \geq 5.2$ detection significance in the SSDF. Each data point corresponds to one cluster candidate defined in §3.1, color-coded by detection significance. The grey background shows the positions of all galaxies in the SSDF field that meet the color criteria described in the text. Note that the white gaps in the background map are due to masked areas surrounding bright stars.

Fig. 3.— The evolution with redshift of SuperCOSMOS $I$-band magnitude (top panel), $[3.6] - [4.5]$ IRAC color (middle panel) and $[4.5]$ IRAC magnitude (bottom panel) for Bruzual & Charlot (2003) simple stellar population models formed at formation redshifts, $z_f$, of 5 (black dot-dashed line). The IRAC color selection criterion we adopt is optimized to find galaxies with $z > 1.3$. The $[3.6] - [4.5]$ color serves as a good redshift indicator for $0.7 < z < 1.5$; above this redshift, the $[3.6] - [4.5]$ color evolution with redshift flattens. To alleviate a known contamination from foreground interlopers at $z \sim 0.3$ (e.g., Papovich 2008; Muzzin et al. 2013) we also apply two magnitude cuts ($[4.5] > 19.5$ and $I > 20.45$).
To identify candidate galaxy clusters, for each galaxy in the SSDF catalog that meets the aforementioned criteria we count the number of similarly selected companions within 1.0′-radius cells. We correct these counts in cells for completeness using the values in Table 5 of Ashby et al. (2013), derived from a Monte Carlo analysis. When creating an IRAC mosaic, the mapping strategy adopted results in some variations in depth across a large field. Therefore, we search for cells with the highest significance overdensity of IRAC-selected sources above the local background. To this aim, for each selected object in the catalog, we also count the completeness-corrected number of similarly selected companions in an outer annulus defined by radii of 5′ and 8′. After taking into account the difference in surface areas, we then derive the excess number of selected companions with respect to the local background, ∆N.

Fig. 2 shows the distribution of completeness-corrected excess number of companions, ∆N, within 1.0′ radius. The mean number of excess galaxies is ⟨∆N⟩ = 0.4. The distribution is skewed toward objects with higher-than-average numbers of companions, suggestive of strong clustering. Similarly to Papovich (2008), we fit a Gaussian to the distribution iteratively clipping at 2σ. The best fit is shown in Fig. 2 and it matches well the low-excess half of the distribution. We find the width of the fitted Gaussian to be σ∆N = 3.7.

The minimum detection significance to adopt, Xf,min, is somewhat arbitrary, trading between completeness and reliability in the derived candidate cluster catalog. We based our conservative choice on extensive tests performed on similar-depth IRAC and SuperCOSMOS data in the Boötes field, where ancillary spectroscopic and much deeper multi-wavelength photometric coverage is also present. We describe these tests in section §3.2, where we estimate that the threshold adopted here ensures a purity of ~80%. Thus we define our sample of candidate clusters of galaxies to be those with ∆N ≥ ⟨∆N⟩ + Xf,min · σ∆N, with Xf,min = 5.2. For the SSDF field this corresponds to objects with ∆N ≥ 19.5 excess companions with respect to the local background (indicated by the vertical dot-dashed red line of Fig. 2).

By design, many of the objects in overdense regions will be identified as companions to multiple sources by our algorithm. To address this, we adopt a similar approach as done in Papovich (2008). We merge cluster candidates applying a friends-of-friends algorithm with a linking length of 1′.0. This results in a final catalog containing Nobs = 279 candidate clusters; their spatial distribution and color-coded detection significance are shown in Fig. 3. We note that none of our clusters is found in the list of SPT SZ-detected clusters reported in Reichardt et al. (2013). This is not unexpected, since the two samples do not overlap in mass-redshift space: the highest-redshift cluster in Reichardt et al. (2013) is at z = 1.075, below the range to which our cluster finding algorithm is sensitive, while the typical mass of the clusters we find is below the mass threshold of the Reichardt et al. (2013) catalog.

In Table 1 we present the list of the top-10 most significant cluster candidates in the SSDF field. Figures 4 and 5 present their images and color-magnitude diagrams.

### 3.2. Purity of the SSDF Cluster Candidates

For a large sample of candidate clusters to be useful for galaxy evolution and cosmological studies, it is important to determine its purity, fpure as a function of the detection significance Xf, defined as

\[
f_{\text{pure}}(X_f) = \frac{N_{\text{real}}}{N_{\text{tot}}} = 1 - \frac{N_{\text{false}}}{N_{\text{tot}}},
\]

where N_{tot} is the total number of clusters candidates above the detection threshold X_f, N_{real} is the number of candidates corresponding to real clusters and N_{false} is the number of false detections. To this aim, we run our cluster finding algorithm on comparable-depth observations from the IRAC Shallow Cluster Survey (ISCS) Eisenhardt et al. (2004).

The ISCS is a wide-field IR-selected galaxy cluster survey carried out using 90s Spitzer/IRAC imaging of the 8.5 deg² Boötes field of the NOAO Deep, Wide-Field Survey (NDWFS, Jannuzi & Dev. 1999). SuperCOSMOS data are available in Boötes of comparable depth to the data available in the SSDF field. Over the past decade the Boötes cluster candidates have been the target of extensive ground- and space-based spectroscopic and photometric campaigns (e.g., Eisenhardt et al. 2008 Ashby et al. 2009, Stern et al. 2010, Brodwin et al. 2011, Stanford et al. 2012, Zeimann et al. 2012, Gonzalez et al. 2012, Brodwin et al. 2013). Using a wavelet search algorithm, operating on photometric-redshift probability distribution functions, described in Brodwin et al. (2006), more than 100 rich cluster candidates at z > 1 were identified. To date 18 of these have been spectroscopically confirmed out to z = 1.9. Note that very accurate photometric redshifts measurements are based on data from the Spitzer Deep, Wide-Field Survey (SDWFS; Ashby et al. 2009) (4 × deeper than the IRAC data used here), deep optical imaging in B_N, R, I bands from the NDWFS, and NIR photometry from the FLAMEX survey Elston et al. (2006) in the J- and K_s-bands. Note, however, that the ISCS is on-going, with many cluster candidates still awaiting confirmation. In the absence of a complete, spectroscopically confirmed catalog of high-redshift clusters, we deem the ISCS the best available survey to test our algorithm as it contains the largest sample (> 10) of spectroscopically confirmed clusters at z > 1.3.

We find the purity of our sample, f_{pure}, to be a monotonic function of X_f reaching f_{pure} = 0.8 for X_{f,min} = 5.2. At this very high detection significance, our algo-
Fig. 4.—Spitzer/IRAC 4.5µm images (4′ × 4′) of the 10 most significant high-redshift candidate galaxy clusters in the SSDF survey. The yellow circles indicate candidate cluster members. The red circle has a radius of 1.0′ that corresponds to an angular diameter distance of ∼0.5 Mpc at z = 1.5.
Fig. 5.— [3.6] − [4.5] color vs. [3.6] magnitude for the 10 most significant high-redshift candidate galaxy clusters in the SSDF survey. The solid lines indicate the median [3.6] − [4.5] color for member galaxies with [4.5] < 21. The dashed lines indicate ±σ, the standard deviation of the color distribution.
4. CLUSTERING OF HIGH REDSHIFT CLUSTERS

Clusters of galaxies reside in the largest dark matter halos, representing higher density peaks of the mass distribution as redshift increases (Springel et al. 2005). Thus the number density of galaxy clusters with redshift is sensitive to the cosmic matter density, \( \Omega_m \), and its evolution with redshift provides constraints on cosmological parameters (e.g., Kitayama & Suto 1996; Wang & Steinhardt 1998).

Several observational studies have previously demonstrated that cluster samples show strong clustering (e.g., Bahcall 1988; Huchra et al. 1990; Postman et al. 1992; Borgani et al. 1993; Gonzalez et al. 2002; Bahcall et al. 2003; Brodwin et al. 2007; Papovich 2008) as measured by their autocorrelation function (Peebles 1980).

High-resolution simulations predict that the cluster correlation function strength increases with redshift for a given mass limit. That is, high-redshift clusters are more strongly clustered, on a comoving scale, than low-redshift clusters of the same mass (Bahcall et al. 2003). Moreover, a relative constancy of the cluster correlation function strength for the most massive clusters at every epoch is predicted: the \( N \) most massive clusters at one epoch should have similar clustering to the \( N \) most massive clusters at a later epoch (Younger et al. 2003).

Currently, both the measured correlation function scale lengths and number densities of several galaxy cluster samples support predictions from standard cold dark matter models (ACDM) (e.g., Abadi et al. 1998; Croft et al. 1997; Peacock & West 1992; Bahcall et al. 1993; Collins et al. 2000; Bahcall et al. 2003; Brodwin et al. 2007; Papovich 2008; Vikhlinin et al. 2009; Hasselfield et al. 2013). Next, we measure these quantities for our cluster sample in order to check our consistency with previous observational studies and model predictions.

4.1. Comoving number density

In order to derive the comoving number density of our cluster sample, we first need to infer the expected redshift distribution \( \phi(z) \) of our sample. To this aim, we employ the COSMOS/UltraVista photometric redshift catalog from Muzzin et al. (2013). This catalog includes 3.6\( \mu \)m and 4.5\( \mu \)m photometry reaching 2 magnitudes deeper than the SSDF. The central panel in Figure 7 shows the smoothed [3.6]–[4.5] color versus redshift distribution of the COSMOS catalog; details on this filtering procedure can be found in Muzzin et al. (2013). The white line marks the [3.6]–[4.5] > –0.1 selection, and the magenta curve follows the peak of the full galaxy distribution. We parametrize this curve with redshift and denote...
it as $S(z)$. At a given redshift, we expect the color distributions of cluster and field galaxies to have the same centroid (given by $S(z)$), but with a smaller scatter for the clusters given their more homogeneous star formation histories and faster evolution (Rettura et al. 2010). As pointed out in §3.1, a significant contamination by $z \sim 0.3$ occurs with a simple IRAC color cut. As we remove this contamination with optical and $[4 \mu m]$ magnitude cuts, in the following analysis we assume that all our clusters are at $z > 1$.

To analyze the colors of our cluster candidates, we consider the 10 brightest galaxies in $[4 \mu m]$ from each of them (the total number of member candidates is $\sim 40$ per cluster). The goal of this selection is two-fold: it removes faint galaxies which are more likely to be contaminants due to photometric errors, and keeps the most massive galaxies which are more likely to share a similar star formation history. The typical scatter in color in clusters after applying this selection is $\sim 0.12$ mag. In comparison, field galaxies at a given redshift have a much larger color scatter of $\sim 0.25$ mag, confirming our expectation that the cluster members are a more homogeneous population.

At this point, we have established that our clusters are likely to follow $S(z)$. In order to derive the cluster redshift distribution, we populate this curve by matching the distribution of mean colors from our full cluster sample. For a given cluster, we define the probability density function of having a mean color $c$ as a normal distribution centered at $\bar{c}$ and standard deviation equal to the standard error of the colors from the 10 brightest cluster members. Then, the full distribution of mean colors, $H(c)$, is given by the sum of the probability density functions from all clusters. This is shown in the right panel of Figure 7. The redshift distribution is then derived as

$$\phi(z) dz = H(S(z)) \frac{dS(z)}{dz} dz,$$

which is displayed in the upper panel of the same figure.

We can calculate the spatial number density of the cluster sample at the pivot redshift $z_p \equiv 1.5$ by combining the redshift distribution $\phi(z)$, the SSDF survey area and the number of cluster candidates, $N_{\text{obs}}$. The number of clusters within $z_p \pm \delta z/2$ can be written as

$$\Delta N = N_{\text{obs}} \frac{\phi(z_p)}{\int \phi(z') dz'} \delta z.$$  \hspace{1cm} (3)

The sampled volume reads as

$$\Delta V = \frac{dV(z_p)}{dz} \delta z = \frac{c \Omega \lambda^2(z_p)}{H(z_p)} \delta z,$$  \hspace{1cm} (4)

where $\chi(z)$ is the comoving radial distance, $H(z)$ is the Hubble function, $c$ is the speed of light and $\Omega = 0.0271$ steradians is the solid angle subtended by the SSDF survey. Hence, the number density for our SSDF cluster sample, $n_c$, at $z = 1.5$ is

$$n_c = \frac{\Delta N}{\Delta V} = (7.6 \pm 0.6) \times 10^{-7} h^3 \text{Mpc}^{-3}.$$  \hspace{1cm} (5)

4.2. Halo model

Our infrared selection of galaxies and the ranking system we adopt for our cluster search algorithm is expected to produce a nearly mass-limited cluster sample at $z > 1.3$. Thus, we assume that our cluster sample is comprised of most dark matter haloes above some characteristic minimum mass $M_{\text{min}}$, following the halo occupation framework Ma & Fry 2003, Seljak 2000, Peacock & Smith 2000, Scoccimarro et al. 2001, Cooray & Sheth 2002, Berlind & Weinberg 2002, Berlind et al. 2003, Kravtsov et al. 2004, Zheng et al. 2005. We model the probability of a halo of mass $M$ to be part of the cluster sample as:

$$N_c(M) = \begin{cases} 1 & \text{if } M \geq M_{\text{min}} \\ 0 & \text{if } M < M_{\text{min}} \end{cases}$$  \hspace{1cm} (6)

Then, the comoving number density of clusters is determined as:

$$n_c = \int_{M_{\text{low}}}^{M_{\text{high}}} \frac{dM}{dM} N_c(M) \frac{M}{M_{\text{min}}}$$  \hspace{1cm} (7)

where $\frac{dM}{dM} (M)$ is the halo mass function from Tinker et al. (2010) and the integration limits are $M_{\text{low}} = 10^{11} M_{\odot}$ and $M_{\text{high}} = 10^{16} M_{\odot}$. In addition, we can calculate the mean mass of the cluster sample as an average over the occupied halos:

$$M_{\text{mean}} = \frac{1}{n_c} \int dM \frac{dn(M)}{dM} N_c(M|M_{\text{min}}) M.$$  \hspace{1cm} (8)

4.3. Angular Correlation Function

An additional observable that links the distribution of dark matter haloes and our clusters is the measurement of their clustering. In particular, we focus on the modeling of the two-point spatial correlation function (SCF, formally $\xi(r)$, Peebles 1980), which represents the excess probability of finding two objects at a separation $r$ with respect to the case of a randomly distributed sample. The formation of dark matter haloes of a given mass $M$ is directly correlated with the dark matter overdensity at their location (Kaiser 1984, Fry & Gaztanaga 1993).

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**Figure 7.** Main: $[3.6] - [4.5]$ color versus redshift distribution of the COSMOS galaxy catalog. In the SSDF we select galaxies with $[3.6] - [4.5] > -0.1$, marked as the horizontal white line. The magenta curve follows the peak of the distribution, where cluster colors will most likely be centered at. **Right:** Distribution of mean colors in our cluster sample. The goal of this selection is two-fold: it removes faint galaxies which are more likely to be contaminants due to photometric errors, and keeps the most massive galaxies which are more likely to share a similar star formation history. The typical scatter in color in clusters after applying this selection is $\sim 0.12$ mag. In comparison, field galaxies at a given redshift have a much larger color scatter of $\sim 0.25$ mag, confirming our expectation that the cluster members are a more homogeneous population.

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Mo & White [1996]. This translates into a scaling between the dark matter and halo SCFs, which is called the halo bias $b_h$:

$$\xi_h(M, r, z) = \xi_{dm}(r, z) b^2_h(M, z).$$  \hspace{1cm} (9)

Our definition of halo mass is that enclosed by a sphere with a density 200 times larger than the critical density of the Universe. We obtain the dark matter correlation function from the CAMB software (Lewis & Bridle 2002), and the halo bias from Tinker et al. (2005), using the updated parameters from Tinker et al. (2005). The bias of the clusters with respect to dark matter then reads

$$b_c = \frac{1}{n_c} \int dM \frac{dn(M)}{dM} N_c(M|\omega_{\min})b_h(M).$$  \hspace{1cm} (10)

Since the observed configuration space of our cluster sample is the celestial sphere, we measure its clustering via the angular correlation function (ACF, formally $\omega(\theta)$). The ACF can be considered as the radial projection of the SCF, which can be computed with the knowledge of the sample’s redshift distribution (Limber 1953; Phillips et al. 1978):

$$\omega(\theta) = \frac{2}{c} \int_0^\infty dz H(z) \phi^2(z) \int_0^\infty dy \xi_c(r = \sqrt{y^2 + D_c^2(z)} \theta^2),$$  \hspace{1cm} (11)

where $\phi(z)$ is the normalized redshift distribution, $D_c(z)$ is the radial comoving distance, $c$ is the speed of light and $\theta$ is the angular separation given in radians. We measure $\omega(\theta)$ with the estimator presented in Hamilton (1993), which counts the number of galaxy pairs with respect to those of a random sample distributed in the same geometry:

$$\hat{\omega}(\theta) = \frac{RR(\theta)GG(\theta)}{(GR)^2(\theta)} - 1,$$  \hspace{1cm} (12)

where $GG$, $GR$ and $RR$ are total number of galaxy-galaxy, galaxy-random and random-random pairs separated by an angle $\theta$. There is no need to include a correction for the integral constraint (Peebles 1980), since it was shown in Martinez-Manso et al. (2014) that it is negligible for this very wide-field survey geometry. We estimate $\omega(\theta)$ errors using jackknife resampling. For this, the entire sample is divided into $N_{\text{jack}} = 32$ spatial regions of equal size. Then, the correlation is run $N_{\text{jack}}$ times, each one excluding one of those regions from the sample. The value of the estimator is the average $\hat{\omega}(\theta)$ of those iterations and the covariance between angular bins is given by

$$C_{jk} = \frac{N-1}{N} \sum_{i=0}^N [\hat{\omega}_i(\theta_j) - \hat{\omega}(\theta_j)] [\hat{\omega}_i(\theta_k) - \hat{\omega}(\theta_k)].$$  \hspace{1cm} (13)

The observed $\omega(\theta)$ function is shown in Figure 8 where the error bars are derived from the diagonal elements of the covariance matrix.

4.3.1. Results

The halo model of the clustering described so far can be fully specified either by fixing $n_c$ through the observed cluster counts (Equation 9), or by fixing $b_c$ through the amplitude of the observed clustering. We refer to these two approaches as the density and clustering fits, respectively. Thus, comparing these sets of results is an excellent way to test the consistency of the model and observations. Our fits are based on the maximization of the likelihood of the model given the data, $L(\text{model}|\text{data}) = e^{-\chi^2/2}$. For the clustering fit, this is specified by:

$$\chi^2 = \sum_{i=0}^{K} \sum_{j=0}^{K} [\omega_m(\theta_j) - \omega(\theta_j)] C_{ij}^{-1} [\omega_m(\theta_k) - \omega(\theta_k)].$$  \hspace{1cm} (14)

Here, $\omega_m$ and $\omega$ are the predicted and observed ACFs (Equations 11 and 12, respectively), $C_{ij}$ is the covariance matrix from Equation 13 and $K = 4$ is the number of angular bins.

For the density fit,

$$\chi^2 = (n_{\text{mod}} - n_{\text{data}})^2 / \sigma^2_{n_{\text{dev}}}.$$

where the variance $\sigma^2_{n_{\text{dev}}}$ is derived from Poisson statistics in the cluster counts. Results are shown in Table 2. We have included the calculation of $r_0$, which marks the distance scale length where $\xi_c(r_0) = 1$, with $\xi_c = \xi_{dm}(z = 0) b^2_c$ for both methods.

Overall, there is a reasonable agreement between the density and clustering sets, differing by less than 2 $\sigma$. Throughout the paper, we make the conservative choice of adopting the clustering sets as fiducial values since they have the larger errors. The two angular correlation functions, $\omega(\theta)$, one corresponding to the prediction based on the number density (solid curve) and one based on the fit of the measured ACF points (dashed curve) are displayed in Figure 8 and found to be consistent within the errors.

The values of $n_c$ and $r_0$ found for our sample of candidate clusters are consistent with those found by previous observational studies in the literature (Adami et al. 1998; Collins et al. 2000; Bahcall et al. 2003; Brodwin et al. 2007, see also Fig. 9 of Papovich 2008) and are well in agreement with predictions based on $\Lambda$CDM cosmological models (Springel et al. 2005). This analysis lends further confidence in the effectiveness of our cluster finding algorithm.

5. SUMMARY

<table>
<thead>
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<th>TABLE 2</th>
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<tr>
<td>Density</td>
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<tr>
<td>$n_c$ (10$^{-7}$h$^3$Mpc$^{-3}$)</td>
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<tr>
<td>$M_{\text{min}}$ (10$^{14}$h$^{-1}$M$_\odot$)</td>
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<td>$M_{\text{mean}}$ (10$^{14}$h$^{-1}$M$_\odot$)</td>
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<tr>
<td>$b_c$</td>
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<tr>
<td>$r_0$ (h$^{-1}$Mpc)</td>
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</table>
We have identified a large sample of massive high-redshift galaxy cluster candidates of galaxies at \( z > 1.3 \) over the 94 deg\(^2\) Spitzer survey of the SPTpol field. Our algorithm identifies the most significant overdensities of galaxies based upon their IRAC color ([3.6] − [4.5] > −0.1), their 4.5\,µm magnitude ([4.5] > 19.5) and requiring non-detection in the shallow SuperCOSMOS I-band data (\( I > 20.45 \)). We identify 279 distant cluster candidates using a \( X_f \geq 5.2 \) detection significance, for which we estimate a \( \sim 80\% \) purity by running our algorithm on comparable-depth observations of the Spitzer surveys of the Boötes field (Eisenhardt et al.

We find that the SSDF cluster sample shows strong clustering. From the angular correlation analysis, we find our sample has a comoving number density \( n_c = (0.7^{+0.3}_{-0.1}) \times 10^{-7}\, h^{-3}\,\text{Mpc}^{-3} \) and a spatial clustering correlation scale length \( r_0 = (32 \pm 7)\, h^{-1}\,\text{Mpc} \). These values are consistent with previous observational studies and match expectations based on ΛCDM high-resolution simulations. The high-redshift cluster sample presented here has a mean mass \( M_{\text{mean}} = 1.7^{+0.8}_{-0.5} \times 10^{14}\,h^{-1}\,M_\odot \). Assuming these clusters grow according to predictions of ΛCDM (e.g., Fakhouri et al. 2010), they will evolve into massive clusters (> \( 5 \times 10^{14}\,h^{-1}\,M_\odot \)) at \( z = 0.2 \).

This study showcases the impact that large Warm Spitzer surveys can have on the identification of large samples of massive clusters of galaxies at very high redshifts in the upcoming years. In particular, this sample has been selected in an area where deep observations for the SZ effect with the SPTpol camera are underway and part of this field has also \( XMM-Newton \) deep X-ray observations from the XXL Survey. These ancillary data will allow us to determine cluster masses for our sample, enabling systematic study of the cluster population in a crucial epoch for their assembly.

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Facilities: Spitzer, UKST

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Fig. 8.—Observed clustering measured via the ACF (points), fitted \( \omega(\theta) \) (dashed curve) and predicted \( \omega(\theta) \) based on the observed number density (solid curve).

The high-redshift cluster sample presented here has a mean mass \( M_{\text{mean}} = 1.7^{+0.8}_{-0.5} \times 10^{14}\,h^{-1}\,M_\odot \). Assuming these clusters grow according to predictions of ΛCDM (e.g., Fakhouri et al. 2010), they will evolve into massive clusters (> \( 5 \times 10^{14}\,h^{-1}\,M_\odot \)) at \( z = 0.2 \).
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