

# Supporting Information

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### Estimation of $\lambda$ and $\alpha(0)$

Here we provide a summary of the procedure described in ref. 16 to calculate  $\lambda$  and  $\alpha(0)$  from available free rotation, stalling, and ensemble biochemical experimental data. In equation 12 of ref. 16  $\alpha(0)$  [i.e.,  $\alpha(\theta)$  at  $\theta=0$ ], and in equation 18  $\lambda$  is given as a function of three energy quantities, shown in bold in the energy diagram in Fig. S1: the work term for the weak binding in the reactant (R) state,  $W^r$ , the free energy barrier in the collision-based theory of binding,  $\Delta G_0^\ddagger$ , and the standard free energy of reaction defined relative to the R state,  $\Delta G_0^0$ . These quantities involved in the binding of ATP are defined relative to the states given in Fig. S1, a state (S) in which the ATP molecule is found in the solution, a reactant state (R), which involves the weak binding of ATP at the entrance site, the transition state (TS), and a product state (P) in which the ATP occupies the binding pocket. We now summarize the procedure to estimate the value of the three quantities, also given in Fig. S1.

- i) The  $W^r$  is estimated to be  $-9.1$  kcal/mol for ATP binding, calculated from the binding affinity of  $-6.3$  kcal/mol to the empty open-conformation subunit [from a dissociation constant (25) of  $25 \mu\text{M}$ ] minus an entropic contribution to the free energy change  $kT \ln kT/hZ = 2.8$  kcal/mol.
- ii) A  $\Delta G_0^\ddagger = 5.0$  kcal/mol estimate is yielded by collision theory, assuming a collision frequency of  $Z = 10^{11} \text{ M}^{-1}\text{s}^{-1}$ . The free rotation rate constant for ATP binding is  $k_{f0} = 2 \times 10^7 \text{ M}^{-1}\text{s}^{-1}$  (1), which is used in equation 8 of ref. 16 to give the above estimate for  $\Delta G_0^\ddagger$ .
- iii) The  $\Delta G_0^0 = -6.0$  kcal/mol estimate for ATP binding can be calculated, according to Fig. S1, by subtracting the binding affinity of  $-6.3$  kcal/mol for the R state from the standard free energy of reaction defined relative to the solution state (S),  $\Delta G_{S0}^0 = -kT \ln K_0 = -12.3$  kcal/mol (3). This method, based on Fig. S1 for evaluating  $\Delta G_0^0$ , is simpler than that used in ref. 16.

### Correcting the Reported Controlled Rotation Data

Here we describe a theoretical method to calculate corrections for the data reported by Adachi et al. (4) due to the effects of (i) missed events and (ii) replacing  $T_0$  with  $T$ .

If  $\theta_j$  is the rotor angle at the center of the  $j^{\text{th}}$  interval, then, according to Adachi et al. (4), we use for the reported (“rep”) rate constant estimates in the 36 intervals the values used in their analysis,

$$k_f^{\text{rep}}(\theta_j) = N(0 \rightarrow 1, j) / T, \quad [\text{S1}]$$

$$k_b^{\text{rep}}(\theta_j) = N(1 \rightarrow 0, j) / T_1(j), \quad [\text{S2}]$$

where  $N(0 \rightarrow 1, j)$  and  $N(1 \rightarrow 0, j)$  are the number of  $0 \rightarrow 1$  and  $1 \rightarrow 0$  events in the interval  $j$ , counted over the whole trajectory ( $j$  has a periodicity of 36). The reported equilibrium constant is then calculated as  $K^{\text{rep}} = k_f^{\text{rep}} / k_b^{\text{rep}}$ . The rate constants estimated using this procedure, obtained by digitizing figure 5 of ref. 4, are reproduced in Fig. S2. Interestingly, the rates for Cy3-ATP binding in the presence and absence of  $\text{P}_i$  are indistinguishable from each other within the scatter in the data, as seen in Fig. S2. So, we treat them as essentially identical and compare them later with the same theoretical prediction. A similar remark applies to Cy3-ADP in Fig. S2.

The reported uncorrected experimental rate estimates defined in Eqs. S1 and S2 are calculated, for sufficiently long trajectories, from the averaged values,

$$k_f^{\text{rep}}(\theta_j) = N(0 \rightarrow 1, j) / T = \frac{\langle N(0 \rightarrow 1) \rangle_j}{t}, \quad [\text{S3}]$$

$$k_b^{\text{rep}}(\theta_j) = \frac{\langle N(1 \rightarrow 0) \rangle_j}{\langle t(\sigma=1) \rangle_j}, \quad [\text{S4}]$$

where the average  $\langle \cdot \rangle_j$  is calculated over a single interval  $j$ , and all quantities depend on  $\tau$ ,  $t$ ,  $k_f$ , and  $k_b$ . The  $k_f$  and  $k_b$  are the actual  $\theta_j$ -dependent rate constants, the subject of the present theoretical predictions. In Eq. S3, the error due to using  $T$  instead of  $T_0$  is also taken into account in the theoretical calculation of uncorrected forward rates, following the procedure used for the experimental estimation of  $k_f^{\text{rep}}$ .

Assuming steady-state conditions, the probability of being in a state  $\sigma = 0$  or 1 at any time is given in terms of the rate constants as

$$p_0 = k_b / (k_f + k_b), \quad p_1 = k_f / (k_f + k_b). \quad [\text{S5}]$$

The survival probabilities in states 0 and 1, denoted by  $S_0$  and  $S_1$ , are both exponential,  $S_0(t) = \exp(-k_f t)$  and  $S_1(t) = \exp(-k_b t)$ , respectively. The probability distributions that the states have lived a lifetime between 0 and  $t$ , that is, that it decayed before  $t$ , are  $[1 - S_0(t)]$  and  $[1 - S_1(t)]$ . Their time derivatives, namely the associated probability densities of lifetimes, the so-called waiting time distributions  $\rho_0(t)$  and  $\rho_1(t)$ , are given by

$$\rho_0(t) = k_f \exp(-k_f t), \quad \rho_1(t) = k_b \exp(-k_b t). \quad [\text{S6}]$$

The probability densities  $\rho_0$  and  $\rho_1$  describe the statistics of times spent in the 0 and 1 states in the trajectories (Fig. S3). We define the local time  $t$  as the time relative to the time when the rotor angle reaches the beginning of an interval  $j$ . In experiment, the finite acquisition time of video frames and the shot noise limit the effective time resolution. As a consequence, states that seem to be of  $\sigma = 1$  but are shorter than  $\tau = 0.1$  s were discarded in the analysis of Adachi et al. (4). We denote by  $p(\sigma_0, n)$  the probability of a succession of  $n$  events of changed occupancy occurring at times  $\{t_1, t_2, \dots\}$  during any given  $10^\circ$  interval, given the condition of the system's being in occupancy  $\sigma_0$  at  $t=0$ . Similarly, the probability that these  $n$  events are missed is denoted by  $p_{\text{miss}}(\sigma_0, n)$ .

The statistics of binding and release events depend on parameters  $\tau$  and  $t$ . In our calculations,  $\tau$  and  $t$  are formally treated as variables and are subsequently assigned the constant values used by the experimentalists, as indicated earlier, so  $p(\sigma_0, n)$  and  $p_{\text{miss}}(\sigma_0, n)$  are probabilities that depend on these variables  $\tau$  and  $t$ . These probabilities can be cast as expectation values with respect to  $\tau_n$  s (defined as  $t_n - t_{n-1}$  in Fig. S3). As an example, for  $n=2$  and  $\sigma_0=0$ ,  $p(0, 2) = p_0 \langle \Theta(t_2 < t < t_3) \rangle$  and  $p_{\text{miss}}(0, 2) = p_0 \langle \Theta(t_2 < t < t_3) \Theta(\tau_2 < \tau) \rangle$ . Here,  $\Theta$  is a generalized Heaviside function which is 1 if the argument is true and 0 otherwise, and  $\langle \cdot \rangle$  is a notation for the integration of the enclosed argument with respect to all  $\tau_n$  s.

In addition to events of the type  $p_{\text{miss}}(\sigma_0, 2)$ , missed “cross-boundary events,” denoted by  $p'_{\text{miss}}(\sigma_0, 2)$ , also have a contribution. Missed cross-boundary events are transitions occurring in a bin with the previous or subsequent transition occurring in another interval, and which transitions are missed due to the time

spent between the two events is too short to be resolved. For example, one such event is when a  $0 \rightarrow 1$  transition occurs during the  $j^{\text{th}}$  bin, and the subsequent transition occurs in a subsequent interval, but the time spent in state 1 is too short (i.e.,  $< \tau$ ) to be detected. Such  $p'_{\text{miss}}(\sigma_0, 1)$ , formally written as  $p_0(\Theta(t_1 < t < t_2)\Theta(\tau_2 < \tau))$ , leads to a convolution integrals. For example, for  $\sigma_0 = 0$ , the time-domain expression is  $p(0,2|t) = p_0 \int_0^\infty \int_0^\infty \int_0^\infty \rho_0(\tau_1)\rho_1(\tau_2)\rho_0(\tau_3)\delta(t - \tau_1 - \tau_2 - \tau_3)d\tau_3d\tau_2d\tau_1$  and  $p_{\text{miss}}(0,2|t) = p_0 \int_0^\infty \int_0^\tau \int_0^\infty \rho_0(\tau_1)\rho_1(\tau_2)\rho_0(\tau_3)\delta(t - \tau_1 - \tau_2 - \tau_3)d\tau_3d\tau_2d\tau_1$  and a Laplace transform ( $t \rightarrow s, \tau \rightarrow u$ ) formally facilitates their evaluation. For  $p(\sigma_0, n)$  the Laplace transform  $t \rightarrow s$  is a formally simple expression (we use a tilde to denote the transform),

$$\begin{aligned} \tilde{p}(\sigma_0, n|s) &= \int_0^t p(\sigma_0, n|t)e^{-st} dt \\ &= p(\sigma_0) \left\langle \exp(-st_n) \frac{1 - \exp(-s\tau_{n+1})}{s} \right\rangle. \end{aligned} \quad [\text{S7}]$$

Using  $t_n = \tau_1 + \dots + \tau_n$ , for  $\sigma_0 = 0$ , an analytic expression follows,

$$\tilde{p}(0, n|s) = p(0) \times \begin{cases} [\tilde{\rho}_0(s)\tilde{\rho}_1(s)]^{n/2}, & n \text{ even,} \\ [\tilde{\rho}_0(s)]^{(n+1)/2} [\tilde{\rho}_1(s)]^{(n-1)/2}, & n \text{ odd,} \end{cases} \quad [\text{S8}]$$

where  $\tilde{\rho}_0(s) = k_f/(s + k_f)$  and  $\tilde{\rho}_1(s) = k_b/(s + k_b)$ . An expression for  $p(1, n|s)$  results by analogy from Eq. S8 (i.e., by interchanging states 0 and 1).

In the calculation of the probability of the missed events, for each short event an additional condition applies, which is formally achieved by introducing an additional variable. For the leading terms of  $n = 2$ , which provide the most contribution, a single additional variable is needed ( $\tau \rightarrow u$ ), leading to a double Laplace transform,

$$\tilde{p}_{\text{miss}}(0, 2|u, s) = p_0 \tilde{\rho}_0(s) \tilde{\rho}_1(s) u^{-1} (s + k_f)^{-1}. \quad [\text{S9}]$$

For the cross-boundary missed events,

$$\tilde{p}'_{\text{miss}}(0, 1|u, s) = p_0 (s u)^{-1} \tilde{\rho}_0(s) [\tilde{\rho}_1(u) - \tilde{\rho}_1(s + u)]. \quad [\text{S10}]$$

Again, by analogy expressions for  $\tilde{p}_{\text{miss}}(1, 2|s)$  and  $\tilde{p}'_{\text{miss}}(1, 1|u, s)$  follow from Eqs. S9 and S10.

The above Laplace-space expressions are then inverted by elementary inversion, facilitated by the use of symbolic computation tools (37). Of special importance are the one-event contributions,

$$p(0,1) = p_0 k_f / (k_f - k_b) [\exp(-k_b t) - \exp(-k_f t)], \quad [\text{S11}]$$

$$p(1,1) = p_1 k_b / (k_b - k_f) [\exp(-k_f t) - \exp(-k_b t)], \quad [\text{S12}]$$

and the two-event contributions,

$$p(0,2) = -\frac{k_f k_b^2}{(k_f + k_b)(k_f - k_b)^2} \times (e^{-k_f t} - e^{-k_b t} + k_f t e^{-k_f t} - k_b t e^{-k_b t}), \quad [\text{S13}]$$

$$p(1,2) = \frac{k_f^2 k_b}{(k_f + k_b)(k_f - k_b)^2} \times (e^{-k_f t} - e^{-k_b t} + k_f t e^{-k_b t} - k_b t e^{-k_b t}). \quad [\text{S14}]$$

The leading terms that effectively determine the probability of missed events within an interval are

$$\begin{aligned} p_{\text{miss}}(0,2) &= k_f k_b^2 e^{-k_f t - k_b 3\tau} (k_f + k_b)^{-1} (k_f - k_b)^{-2} \\ &\quad (e^{k_f 3\tau} - e^{k_b 3\tau} + k_f t e^{k_f 3\tau} - k_f t e^{k_b 3\tau} \\ &\quad - k_b t e^{k_f 3\tau} + k_b t e^{k_b 3\tau} - k_f 3\tau e^{k_f 3\tau} + k_b 3\tau e^{k_f 3\tau}). \end{aligned} \quad [\text{S15}]$$

For the cross-boundary missed events, the Laplace inversion yields for the leading terms

$$\begin{aligned} p'_{\text{miss}}(0,1) + p'_{\text{miss}}(1,1) &= \\ &\quad \left[ e^{-k_f t} e^{-k_b 3\tau} - e^{-k_b 3\tau} - \frac{k_f e^{-k_f t}}{k_f - k_b} + \frac{k_f e^{-k_b t}}{k_f - k_b} \right. \\ &\quad \left. + k_f k_b e^{-k_b 3\tau} \left( \frac{1}{k_f k_b} + \frac{e^{-k_f(t-3\tau)}}{k_f(k_f - k_b)} - \frac{e^{-k_b(t-3\tau)}}{k_b(k_f - k_b)} \right) \right]. \end{aligned} \quad [\text{S16}]$$

To obtain the reported rates from Eqs. S3 and S4 we subtract the probability of missed events from the total probability of events,

$$k_f^{\text{rep}}(\theta) \cong t^{-1} [p(0,1) + p(0,2) + p(1,2) - p_{\text{miss}}(0,2) - p_{\text{miss}}(0,1) - p_{\text{miss}}(1,1)], \quad [\text{S17}]$$

$$\begin{aligned} k_b^{\text{rep}}(\theta) &\cong (tp_0)^{-1} \langle N(1 \rightarrow 0, \tau, t) \rangle_j \\ &= (tp_0)^{-1} [p(1,1) + p(1,2) + p(0,2) \\ &\quad - p_{\text{miss}}(0,2) - p_{\text{miss}}(0,1) - p_{\text{miss}}(1,1)]. \end{aligned} \quad [\text{S18}]$$

In the calculations we neglect higher-order contributions (i.e., terms where  $n \geq 3$ ). This approximation is valid if  $t$  is smaller or comparable with the inverse of the rates, that is,  $t \leq 1/k_f$  and  $t \leq 1/k_b$ . In our calculations it is always verified that this condition is met. We note that the definitions in Eqs. S1 and S2, for sufficiently slow rotation, yield  $k_f^{\text{rep}} < k_b^{\text{rep}}$ , even if the actual rate constants have the opposite relation,  $k_f > k_b$ . Thus, the latter condition could effectively be masked in the experiments, but with the necessary corrections the actual rates can be recovered. As discussed in the text, the number of missed  $0 \rightarrow 1 \rightarrow 0$  events is appreciable at low  $\theta$  values, that is, at  $\theta \sim -50^\circ$ , when  $k_b$  becomes large compared with  $1/t$  and much larger than  $k_f$ , resulting in significant corrections for both  $k_f$  and  $k_b$ , as seen in Fig. S4. The same remark applies to  $0 \rightarrow 1 \rightarrow 0$  events, when  $k_f > 1/t$  and  $k_f \ll k_b$ , but these conditions are not reached in the experiments.







