

An enlarged cosmic shear survey with the William Herschel Telescope

Richard Massey,^{1,2★} Alexandre Refregier,³ David J. Bacon,⁴ Richard Ellis²
and Michael L. Brown⁴

¹*Institute of Astronomy, Madingley Road, Cambridge CB3 0HA*

²*California Institute of Technology, Pasadena, California 91125, USA*

³*Service d'Astrophysique, CEA Saclay, F-91191 Gif sur Yvette, France*

⁴*Institute for Astronomy, Blackford Hill, Edinburgh EH9 3HJ*

Accepted 2005 March 10. Received 2005 February 17; in original form 2004 April 15

ABSTRACT

We report the results of a cosmic shear survey using the 4.2-m William Herschel Telescope on La Palma, to a depth of $R = 25.8$ ($z \approx 0.8$), over 4 deg^2 . The shear correlation functions are measured on scales from 1 to 15 arcmin, and are used to constrain cosmological parameters. We ensure that our measurements are free from instrumental systematic effects by performing a series of tests, including a decomposition of the signal into E - and B -modes. We also reanalyse the data independently, using the shear measurement pipeline developed for the COMBO-17 survey. This confirms our results and also highlights various effects introduced by different implementations of the basic ‘Kaiser–Squires–Broadhurst’ shear measurement method. We find that the normalization of the matter power spectrum on $8 h^{-1} \text{ Mpc}$ scales is $\sigma_8 = (1.02 \pm 0.15)(0.3/\Omega_m)^{1/2}$, where the 68 per cent confidence limit error includes noise, sample variance, covariance between angular scales, systematic effects, redshift uncertainty and marginalization over other parameters. We compare these results with other cosmic shear surveys and with recent constraints from the *Wilkinson Microwave Anisotropy Probe* experiment.

Key words: gravitational lensing – cosmology: observations – large-scale structure of Universe.

1 INTRODUCTION

Weak gravitational lensing by large-scale structure, or ‘cosmic shear’, has emerged as a powerful cosmological probe, as it is directly sensitive to foreground mass (for reviews, see Bernardeau 1999; Mellier 1999; Bartelmann & Schneider 2000; Wittman 2002; Refregier 2003). A measurement of cosmic shear is therefore closely tied to cosmological theories, which are principally concerned with the distribution of dark matter. In particular, the systematic biases of this technique are not limited by unknown physics such as biasing (Dekel & Lahav 1999; Gray et al. 2002; Hoekstra et al. 2002b; Smith et al. 2003a; Weinberg, Davé & Hernquist 2004) or the mass–temperature relation for X-ray-selected galaxy clusters (Pierpaoli, Scott & White 2001; Viana, Nichol & Liddle 2002; Huterer & White 2003).

Cosmic shear surveys are rapidly growing in size and precision (Rhodes, Refregier & Groth 2001; Hoekstra et al. 2002a; Refregier, Rhodes & Groth 2002; Van Waerbeke et al. 2002; Bacon et al. 2003; Brown et al. 2003; Hamana et al. 2003; Jarvis et al. 2003). Cosmological parameter constraints from these surveys are now approaching the precision of other methods.

However, cosmic shear surveys can be subject to several systematic biases of their own. Imperfect telescope tracking, telescope flexure or optical misalignment within the camera, even at a level that is acceptable for most purposes, can artificially distort images in a way that mimics cosmic shear.

The survey described in this paper represents a culmination of effort at the William Herschel Telescope (WHT). We have combined the experience of instrumentalists with detailed image simulations and careful data analysis to control the various sources of systematic error. Our first cosmic shear paper (Bacon, Refregier & Ellis 2000) reported an initial detection of cosmic shear using a 0.5 deg^2 survey with the William Herschel Telescope. The second paper (Bacon et al. 2003) compared the WHT shear signal with an independent measurement using the Keck II telescope, and examined systematic effects from these two very different instruments. In this paper, we extend our WHT survey to cover 4 deg^2 to constrain cosmological parameters, while paying great care in monitoring and correcting systematic effects. We also now test our entire software for shear measurement and cosmological parameter estimation by comparing it to external code, developed independently for the COMBO-17 survey by Brown et al. (2003).

This paper is organized as follows. In Section 2 we describe our survey strategy and observational parameters. In Section 3 we present our results and draw constraints upon cosmological

★E-mail: rjm@astro.caltech.edu

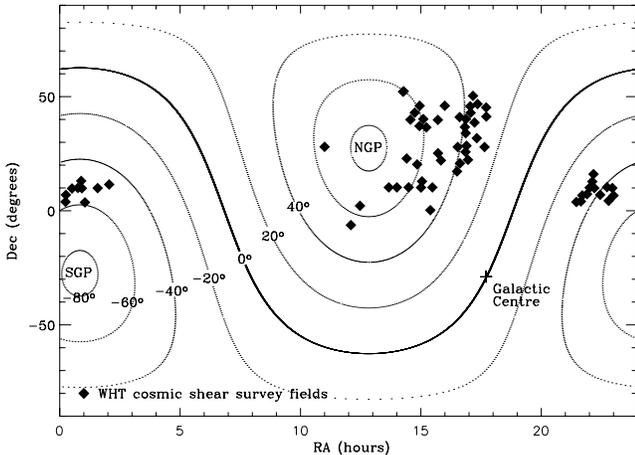


Figure 1. The locations of WHT pointings in our deep pencil-beam survey strategy. The Galactic latitudes were tuned to provide sufficient foreground stars within each image to successfully model and correct for variations in the PSF.

parameters. In Section 4 we test for the absence of any systematic errors. In Section 5, we present a second set of results obtained via an independent shear measurement pipeline. We conclude in Section 6.

2 OBSERVATIONS AND DATA REDUCTION

We have acquired 4 deg² of imaging to $R = 25.8$ (for a point source at 5σ) with the Prime Focus Imaging Camera (PFIC) of the William Herschel Telescope on La Palma. The median seeing was 0.69 arcsec and no exposures had a seeing worse than 1 arcsec. The pixel size was 0.24 arcsec. As shown in Fig. 1, pointings were scattered randomly in a pencil-beam survey between Galactic latitudes of 30° and 70°. This was tuned to provide ~ 1.5 star arcmin⁻², with which we could measure the point spread function (PSF) across each field. The only selection criterion was to avoid foreground stars brighter than $R \approx 11$ in the Digitized Sky Survey or Automated Plate Measuring machine (APM) catalogues.

Cosmic shear statistics have already been presented from the first square degree of this survey in Bacon et al. (2003). That data consisted of eight 8×16 arcmin² images and 11 16×16 arcmin² images taken after the addition of a second, identical CCD to the PFIC. During June and 2002 August, we observed an additional 41 16×16 arcmin² pointings. When combining statistical measurements from all of these fields, we weight the contribution of the large fields to be twice that of the small fields. This scheme is ideal for measurements on small angular scales, where the large fields do indeed contain twice as many pairs of galaxies as the small fields. The scheme is non-optimal in exploiting the extra signal on large scales contained via additional pairs across different CCDs. However, in general, as both large and small fields were randomly deployed, no bias should result from our weighting scheme.

For each survey field we took four 900-s exposures, each dithered by a few arcsec from the previous one. This strategy enabled a continual monitoring of astrometric distortions within the telescope, cosmic ray removal and lower overheads in the event of inclement weather. Data reduction then proceeded for the exposures exactly as in Bacon et al. (2003): see that paper for more details. After bias subtraction and flat-fielding, fringing remained in the R -band images. This could have been prevented by observing at a shorter wavelength, but at a cost to the observed number density of back-

ground sources. To remove this, a fringe frame was compiled from all the exposures in each night. A multiple of this was subtracted from each image which minimized fringing, to a negligible level < 0.05 per cent of the background noise. The four dithers for each field were then realigned (using linear interpolation between adjacent pixels to allow subpixel offsets) and stacked (with 3σ -clipping to remove cosmic rays).

Objects were located on the final images using HFINDPEAKS from the IMCAT package by Kaiser, Squires & Broadhurst (1995, hereafter KSB). Following the recommendations of Massey et al. (2001), objects within 10 arcsec of saturated stars or 5 arcsec of the edge of the CCDs were masked and removed. We also removed noisy objects from the catalogue with cuts in size, signal-to-noise ratio and ellipticity of $r_g > 1$ pixel, $\nu > 15$ and $|\epsilon| < 0.5$. This is the same procedure as that adopted in Bacon et al. (2003), and leaves 15.2 galaxy arcmin⁻² in the final catalogue, with a median magnitude of 23.5 ± 0.2 . According to Cohen et al. (2000) as before, this corresponds to a median source redshift of $z_s \approx 0.80 \pm 0.06$.

The observed ellipticities of galaxies, ϵ , were formed from combinations of their Gaussian-weighted quadrupole moments, and then corrected for convolution with the PSF of the telescope. The shape of the PSF was measured from stars in each image, and then interpolated to the positions of galaxies via a polynomial fit within each CCD separately. The order of this polynomial was varied from field to field: it was occasionally raised to the fourth order but lowered to the second order where possible. This limited spurious power on small scales and particularly around the edges of the field. Even so, an acceptable polynomial fit to the PSF was not always possible. We conservatively discarded 14 fields, all of which were otherwise within specifications, but where the stellar ellipticities remained correlated after correction. More sophisticated PSF fitting methods have recently been suggested by Hoekstra (2004) and Jarvis & Jain (2005). Such techniques may improve the interpolation, and thus increase the number of usable fields, but they have not been investigated here.

The shear susceptibility factor for each galaxy, P^γ , was determined from the higher-order shape moments of the galaxy and the interpolated PSF. The shear susceptibility is a notoriously noisy quantity; in this implementation of KSB, we chose to fit P^γ with a cubic polynomial as a function of galaxy size, r_g . We can then form shear estimators for each galaxy, $\gamma = \epsilon / P^\gamma$. We apply a final calibration factor of $(0.85 \pm 0.04)^{-1}$ to these shears. This calibration factor was determined by Bacon et al. (2001), using our shear measurement pipeline upon simulated WHT images with a known input signal.

Each set of four dithered exposures were also used to continually monitor astrometric distortions within the telescope. As observed in Bacon et al. (2001), these closely follow the engineering predictions in the PFIC manual of $\gamma_{\text{tangential}} = 0$, $\gamma_{\text{radial}} = -8.2 \times 10^{-5} r^2$, where r is measured in arcmin from the field centre. This affects the shear estimate of an average galaxy in a large exposure by less than 2×10^{-3} . A galaxy in the corner of a large exposure is altered by 1×10^{-2} . Since this is no longer an *entirely* negligible effect in our enlarged survey, we subtract this distortion from the final shear catalogues using the shear addition and subtraction operators in Bernstein & Jarvis (2002).

The resulting distribution of shear estimators is shown in Fig. 2. Both components of shear are well fitted by the normalized probability density function

$$P(\gamma_i) \approx 0.26 \frac{\exp(-x^2/0.65^2)}{(x^2 + 0.27^2)}. \quad (1)$$

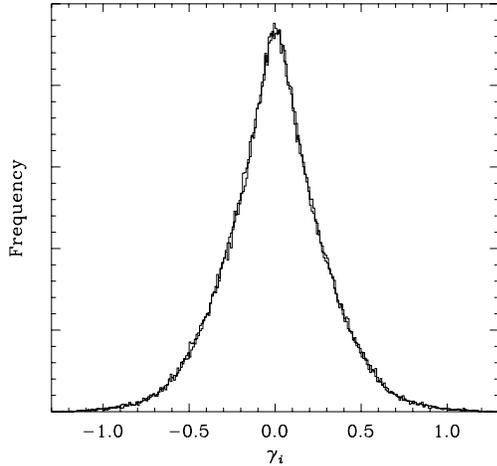


Figure 2. The distribution of measured shear estimators from all galaxies used in the WHT survey. A fitting function is given in the text. The two components of shear are shown superimposed; their distributions are indistinguishable. Note, however, the extended wings typical of the KSB method.

Extended, Cauchy-like wings are a salient feature of the KSB method, where ellipticity measurements are formed from ratios of (noisy) quadrupole moments.

3 RESULTS

3.1 Shear–shear correlation functions

The power spectrum of the weak-lensing shear is given by

$$C_\ell^\gamma = \frac{9}{16} \left(\frac{H_0}{c} \right)^4 \Omega_m^2 \int_0^{\chi_h} \left[\frac{g(\chi)}{D_A(\chi)} \right]^2 P \left(\frac{\ell}{r}, \chi \right) d\chi, \quad (2)$$

where χ is a comoving distance; χ_h is the horizon distance, $D_A(\chi)$ is an angular diameter distance, $g(\chi)$ is the lensing weight function and $P(k, z)$ is the underlying three-dimensional (3D) distribution of mass in the Universe. The two-point shear correlations functions can be expressed in terms of the power spectrum as

$$C_1(\theta) = \frac{1}{4\pi} \int_0^\infty C_\ell^\gamma [J_0(\ell\theta) + J_4(\ell\theta)] \ell d\ell \quad (3)$$

$$C_2(\theta) = \frac{1}{4\pi} \int_0^\infty C_\ell^\gamma [J_0(\ell\theta) - J_4(\ell\theta)] \ell d\ell. \quad (4)$$

These can be measured by averaging over galaxy pairs, as

$$C_1(\theta) = \langle \gamma_1^r(\mathbf{r}) \gamma_1^r(\mathbf{r} + \boldsymbol{\subseteq}) \rangle \quad (5)$$

$$C_2(\theta) = \langle \gamma_2^r(\mathbf{r}) \gamma_2^r(\mathbf{r} + \boldsymbol{\subseteq}) \rangle, \quad (6)$$

where θ is the separation between the galaxies and the superscript r denotes components of shear rotated so that $\gamma_1^r(\gamma_2^r)$ in the first galaxy points along (at 45° from) the vector between the pair. A third shear–shear correlation function can be formed,

$$C_3(\theta) = \langle \gamma_1^r(\mathbf{r}) \gamma_2^r(\mathbf{r} + \boldsymbol{\theta}) \rangle + \langle \gamma_2^r(\mathbf{r}) \gamma_1^r(\mathbf{r} + \boldsymbol{\theta}) \rangle, \quad (7)$$

for which the parity invariance of the Universe requires a zero signal. $C_3(\theta)$ can therefore be used as a first test for the presence of systematic errors in our measurement.

To perform the measurement in practice, we first measure the shear correlation functions in each field, using all galaxy pairs within

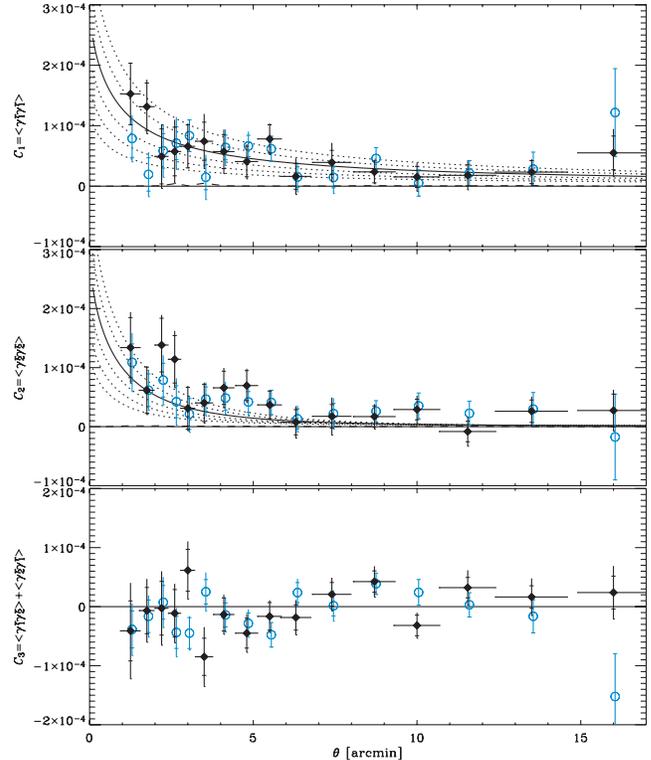


Figure 3. Correlation functions of the shear field measured in our 4 deg² WHT survey. The solid data points show our measurement. The inner error bars are for statistical errors only; the outer error bars also include full non-Gaussian sample variance. The blue circles show measurements of our data by the COMBO-17 pipeline, artificially adjusted to correct for its slightly higher source redshift distribution and to permit an easy comparison. The solid line shows the theoretical prediction for a Λ CDM model with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $\Gamma = 0.21$ and $\sigma_8 = 1.0$, assuming a median source redshift of $z_s = 0.8$ and using the fitting functions of Smith et al. (2003b). The dotted lines show similar theoretical predictions, but with σ_8 ranging from 0.7 (bottom) to 1.2 (top). The dashed lines barely visible above the x axis show the correlation of galaxy shears with the observed PSF anisotropy.

various θ bins. To obtain a combined result for the entire survey, we then average the binned values for each field. We find that the measured correlation functions are quite sensitive to changes in the binning scheme. The highly non-Gaussian distribution of shear estimators (see Fig. 2) means that individual outliers can significantly bias measurements. Both the correlation functions and the subsequent constraints on cosmological parameter constraints can move within their full 1σ statistical error bars. During the averaging, we therefore introduce 3σ -clipping in each bin to remove some outliers, and further 3σ -clipping in $C_3(\theta)$ and the star–galaxy cross-correlation functions $C_1^{\text{SG}}(\theta) + C_2^{\text{SG}}(\theta)$ to eliminate flaws with PSF correction (see Section 4.2). This leaves between $N_f = 40$ and 43 fields used for each angular bin. Our final choice of bin size yields representative central values. The result is shown in Fig. 3. We shall separately estimate the amount of instability caused by the non-Gaussian wings, and increase the error upon cosmological parameter constraints accordingly.

In order to derive any constraints on cosmological parameters, it will also be necessary to know the covariance of $C_i(\theta)$ between different angular bins. Our pencil-beam survey strategy with many independent fields makes it easy to measure their covariance

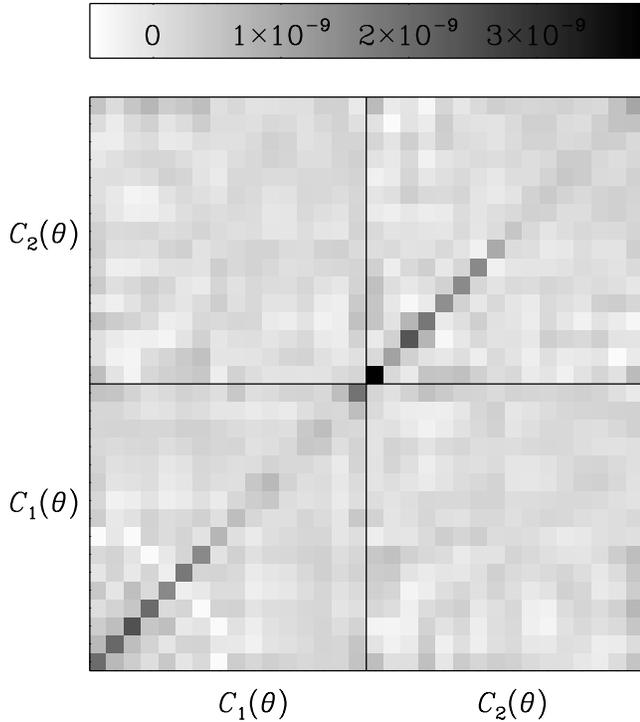


Figure 4. Covariance matrix of different angular bins of the shear correlation functions shown in Fig. 3, from small to large θ again. For example, the bottom-left quarter shows $\text{cov}[C_1(\theta), C_1(\vartheta)]$. The diagonal elements within this are the variances in each bin.

matrix,

$$\text{cov}[C_i(\theta), C_j(\vartheta)] \simeq \frac{1}{N_f^2} \sum_{f=1}^{N_f} [C_i^f(\theta) - C_j(\theta)] [C_i^f(\vartheta) - C_j(\vartheta)], \quad (8)$$

where the summation is over all fields, and the superscript f denotes correlation functions calculated in one field alone. This matrix is depicted in Fig. 4, and shows the significant covariance, especially between adjacent bins.

3.2 Cosmological parameter constraints

We now use a maximum-likelihood method to determine the constraints set by our observations upon the cosmological parameters Ω_m , the total mass-density of the Universe, and σ_8 , the normalization of the matter power spectrum at $8 h^{-1}$ Mpc. The analysis directly uses the observed correlation functions $C_1(\theta)$ and $C_2(\theta)$, proceeding as in Bacon et al. (2003), except that theoretical predictions for the non-linear power spectrum are calculated via the updated fitting functions of Smith et al. (2003b) rather than those by Peacock & Dodds (1996). This has the effect of lowering our final constraint on $\sigma_8 \Omega_m^{1/2}$ by approximately 5 per cent. We use the fitting functions for the linear transfer function suggested by (Bardeen et al. 1986). Note that, although we shall perform an E/B decomposition, we fit C_1 and C_2 rather than the E -mode signal to avoid degeneracies arising from the finite size of our survey. We will use the E/B decomposition as an a posteriori consistency check for systematic effects on relevant scales; because of such contamination, we discarded the first and last data points shown in Fig. 3 (see the discussion in Section 4).

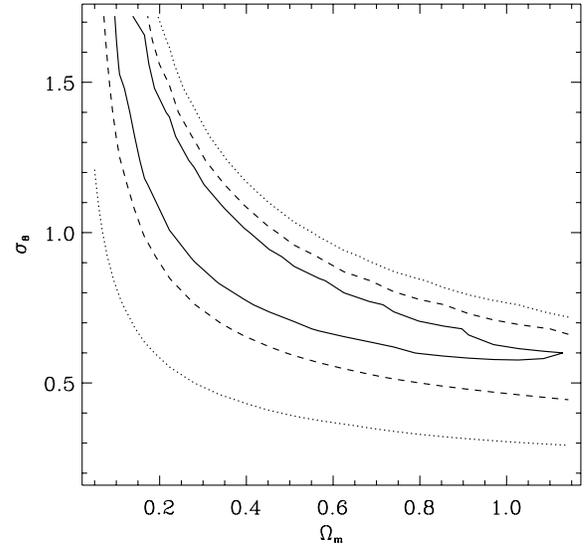


Figure 5. Constraints upon cosmological parameters Ω_m and σ_8 , from a maximum-likelihood analysis of our WHT cosmic shear survey data. The 68.3 per cent (solid), 95.4 per cent (dashed) and 99.7 per cent (dotted) confidence limits include statistical errors and non-Gaussian cosmic variance. However, they include neither the calibration of the shear measurement method, nor uncertainty in the source galaxy redshift distribution. These sources of error are considered separately in the text.

The theoretical correlation functions were first calculated from equation (2) on a 2D grid across the Ω_m versus σ_8 plane. The power spectrum shape parameter was set to $\Gamma = 0.21$, consistent with recent observations of clustering in galaxy redshift surveys (Percival et al. 2001; Szalay et al. 2003). The median redshift for source galaxies was fixed to $z_s = 0.8$ for WHT and $z_s = 1.0$ for Keck. Errors on these parameters will be propagated separately into our final constraints.

We then fitted the observed shear correlation functions $\mathbf{d}(\theta)$ to the theoretical predictions calculated at the centres of each bin $\mathbf{t}(\theta)$, computing the log-likelihood function

$$\chi^2 = [\mathbf{d}(\theta) - \mathbf{t}(\theta, \Omega_m, \sigma_8)]^T \times \{\text{cov}[C_i(\theta), C_j(\vartheta)]\}^{-1} [\mathbf{d}(\vartheta) - \mathbf{t}(\vartheta, \Omega_m, \sigma_8)] \quad (9)$$

throughout the grid. We thus explore parameter space in this plane, and minimize χ^2 to find the best-fitting cosmological model. To compute confidence contours, we numerically integrate the likelihood function

$$L(\Omega_m, \sigma_8) = e^{-\chi^2/2}. \quad (10)$$

Our constraints on cosmological parameters are presented in Fig. 5, and the constraints from our Keck survey (Bacon et al. 2003) are reproduced in the same format in Fig. 6. In both cases, the contours show 68.3, 95.4 and 99.7 per cent confidence limits (CLs), including statistical errors and non-Gaussian sample variance. They reveal the well-known degeneracy between Ω_m and σ_8 when using only two-point statistics.

A good fit to the 68.3 per cent confidence level from our WHT data is given by

$$\sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{0.50} = 1.02 \pm 0.13, \quad (11)$$

while the Keck data are well fitted by

$$\sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{0.52} = 1.01 \pm 0.19. \quad (12)$$

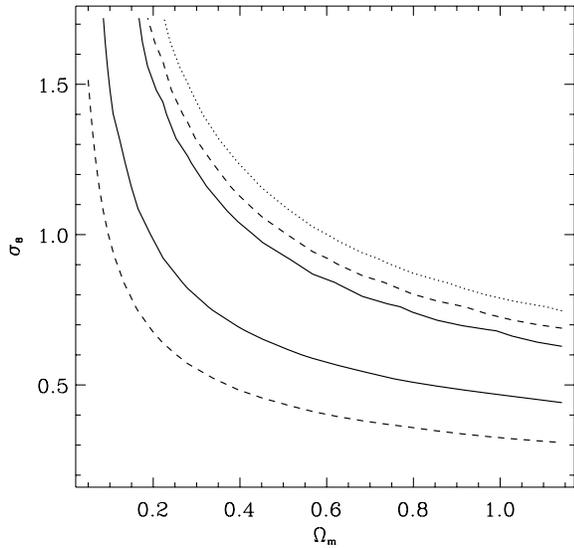


Figure 6. Constraints upon cosmological parameters from the Keck cosmic shear survey by Bacon et al. (2003), showing the 68.3, 95.4 and 99.7 per cent confidence limits as in Fig. 5. Only one edge of the 99.7 per cent confidence contour is visible inside this parameter range.

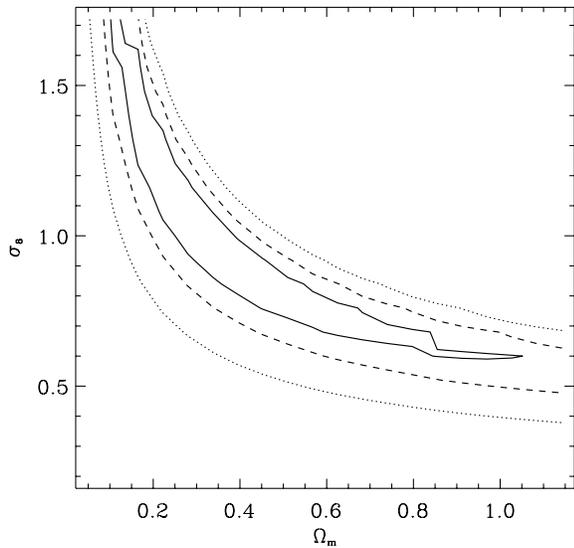


Figure 7. Constraints upon cosmological parameters for the combination of both WHT and Keck surveys, showing the 68.3, 95.4 and 99.7 per cent confidence limits, as in Fig. 5.

The multiplication of the respective likelihood functions provides a constraint from a combined survey. Such confidence contours are shown in Fig. 7, with the 68.3 per cent confidence level well fitted by

$$\sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{0.50} = 1.02 \pm 0.12, \quad (13)$$

for $0.1 < \Omega_m < 0.7$.

Note that all of these constraints include only the statistical error and sample variance. We can propagate other sources of error by noting that

$$C_i(5') \propto \Omega_m^{1.46} \sigma_8^{2.45} z_s^{1.65} \Gamma^{-0.11} (P^i)^{-2}, \quad (14)$$

where $i = 1$ and 2 , and P^i is the shear calibration factor, in a fiducial Λ cold dark matter (Λ CDM) cosmological model with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $\Gamma = 0.21$ and $\sigma_8 = 1.0$. Adding in turn to our constraint (13): a 6 per cent instability due to non-Gaussian errors, a 10 per cent source redshift uncertainty, a 15 per cent prior on Γ and a 5 per cent shear calibration uncertainty, gives a final 68.3 per cent CL constraint for the combined survey of

$$\begin{aligned} \sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{1/2} &= 1.02 \pm 0.12 \pm 0.06 \pm 0.066 \pm 0.006 \pm 0.04 \\ &= 1.029 \pm 0.15, \end{aligned} \quad (15)$$

where the various errors have been combined in quadrature on the second line. This result now includes all contributions to the total error budget: (non-Gaussian) statistical noise, sample variance, covariance between different angular scales, shear calibration error, source redshift uncertainty, and marginalization over Γ .

3.3 Shear variance

For historical reasons, cosmic shear results are often expressed as the variance of the shear field in circular cells on the sky. For a top-hat cell of radius θ , this measure is related to the shear correlation functions by

$$\sigma_\gamma^2 \equiv \langle |\bar{\gamma}|^2 \rangle = \frac{2}{\pi} \int_0^\infty C_\ell^\gamma(\ell) [J_1(\ell\theta)]^2 \ell \, d\ell \quad (16)$$

$$\simeq \frac{2}{\theta^2} \int_0^\theta [C_1(\vartheta) + C_2(\vartheta)] \, d\vartheta, \quad (17)$$

where we have used a small-angle approximation involving the Bessel functions. Note that the shear variance is more strongly correlated on different angular scales in this form than they are as correlation functions.

In practice, data are not available on all the scales necessary to perform this integration. The correlation functions have not been calculated on scales smaller than 1 arcmin and are contaminated by systematic effects on scales smaller than 2 arcmin (see Section 4). We determine the deficit in the measured values as a function of θ by extrapolating the data through these scales using the theoretical predictions given by the best-fitting cosmological model determined in Section 3.2. This deficit ($\sim 2 \times 10^{-5}$ at 3 arcmin and 1×10^{-5} at 5 arcmin) is then added back on to our measured data points.

We present our results as the variance of shear in cells, and compare them with those from similarly deep lensing surveys in Fig. 8. These surveys use data from the 8.5 deg^2 , $z_s = 0.8$ VIRMOS-DESCART survey on the 3.6-m CFHT by Van Waerbeke, Mellier & Hoekstra (2005); the 1.25 deg^2 , $z_s = 0.85$ COMBO-17 survey on the 2.2-m La Silla telescope by Brown et al. (2003); the 0.36 deg^2 , $z_s = 0.9$ Medium Deep Survey (MDS) with the Wide Field and Planetary Camera on the *Hubble Space Telescope* (HST) by Refregier et al. (2002); 0.27 deg^2 of random fields observed to $z_s = 1.0$ in parallel-mode with the Space Telescope Imaging Spectrograph (STIS) on the HST by Rhodes et al. (2005); and the 0.6 deg^2 , $z_s = 1.0$ pencil-beam survey using the 10-m Keck II telescope by Bacon et al. (2003). For ease of comparison between these different surveys, all of the results have been scaled to the values that would have been obtained if their median source redshift had been $z_s = 0.8$, and using the theoretical prediction that $\sigma_\gamma^2 \propto z_s^{1.65}$. Results from deeper surveys are thus shown slightly lower here than in their original papers.

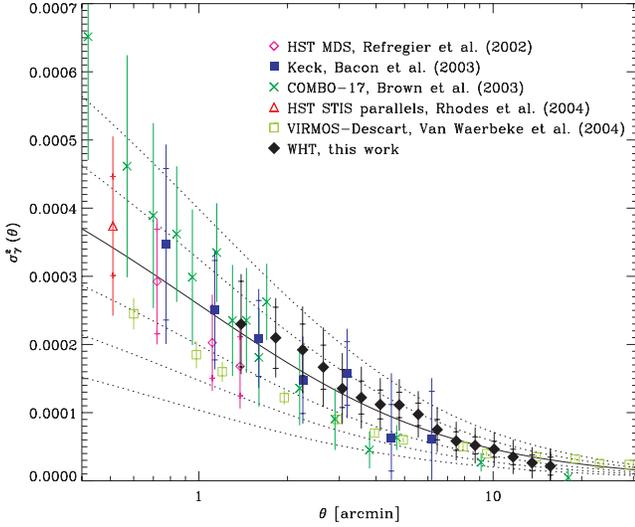


Figure 8. Shear variance in (circular, top-hat) cells, as a function of the radius of the cells. Our results are compared with those from similarly deep surveys by other groups. For ease of comparison, all results have been rescaled as if their median source redshift had been the same (see the text). As in Fig. 3, the theoretical curves are predictions for a Λ CDM cosmology, with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $\Gamma = 0.21$ and σ_8 ranging from 0.7 (bottom) to 1.2 (top).

4 TESTS FOR SYSTEMATIC BIASES

The validity of any cosmic shear result depends sensitively upon the treatment of systematic errors and the control of observational biases. Almost all systematic effects, whether they are due to a background gradient, astrometric distortions within the telescope or imperfectly corrected PSF anisotropy, act to increase the observed correlations between galaxy shapes. All of the effects can mimic cosmic shear and the most important task incumbent upon any weak-lensing survey is to prove that its systematic effects are controlled to a negligible level. As already described in Section 2, the astrometric distortions in WHT have been corrected for, and the basic data reduction was performed sufficiently carefully to eliminate most biases. In this section, we discuss further tests for other sources of residual systematic effects.

4.1 E - B decomposition

The correlation functions can be recast in terms of E (gradient) and B (curl) modes of the shear field (Crittenden et al. 2000; Pen, Van Waerbeke & Mellier 2002). Gravitational lensing is expected to produce only E -modes, except for a very low level of B -modes due to lens-lens coupling along a line of sight (Schneider, van Waerbeke & Mellier 2002). Systematic effects are likely to affect both E - and B -modes equally. The presence of any non-zero B -mode would therefore be a useful indication of contamination from other sources.

E - and B -modes correspond to patterns within an extended region on the sky. The separation cannot be performed locally, but requires the shear correlation functions to be integrated over a wide range of angular scales. In practice, we cannot perform these integrals exactly because our correlation function data extends only between ~ 2 and 16 arcmin. In other words, a shear field measured within a finite aperture cannot be uniquely split into distinct E - and B -mode components: some correction will always be necessary.

We find the frequently used decomposition into aperture mass $M_{\text{ap}}(\theta)$ and $M_{\perp}(\theta)$ -modes (Bartelmann & Schneider 2000) to be particularly unstable with our data. The correlation functions need to be integrated between 0 and 2θ , with a lot of weight placed upon their values at small angular scales. Since these are changing rapidly, the end result becomes even more sensitive to the bin spacing. Furthermore, our measured correlation functions are least reliable at small separations. This is likely to be the case in any real data because of small-scale effects such as overlapping galaxy isophotes. Since data from these small scales need to be included in all subsequent integrals, any bias there would adversely affect the aperture mass on all scales.

We therefore prefer another method for E/B decomposition. Following Crittenden et al. (2000), we calculate

$$C_E(\theta) \equiv C_1(\theta) + 2 \int_{\theta}^{\infty} \left(1 - \frac{3\vartheta^2}{\vartheta^2}\right) \frac{C_1(\vartheta) - C_2(\vartheta)}{\vartheta} d\vartheta, \quad (18)$$

which contains only the E -mode signal and

$$C_B(\theta) \equiv C_2(\theta) - 2 \int_{\theta}^{\infty} \left(1 - \frac{3\vartheta^2}{\vartheta^2}\right) \frac{C_1(\vartheta) - C_2(\vartheta)}{\vartheta} d\vartheta, \quad (19)$$

which contains only the B -mode signal. These can be calculated using data exclusively from scales larger than θ . A correction will still need to be made for absence of data on scales larger than 16 arcmin, but this is in a regime where the expected signal (and the necessary correction) is small. As can be seen from the above equation, a function of θ (not only a constant of integration) must be added to $C_E(\theta)$ and subtracted from $C_B(\theta)$ (cf. Pen et al. 2002). We calculate this function by using theoretical predictions for the best-fitting cosmological model (as determined in Section 3.2) to extrapolate our data to infinity. The size of this correction is approximately one-third of the size of the measured E -mode signal. The correction is 2.4×10^{-5} at 5 arcmin and 1.5×10^{-5} at 10 arcmin.

An E/B decomposition of our data is shown in Fig. 9. On scales of $1.7 < \theta < 15$ arcmin, we find a B -mode signal consistent with zero, confirming the absence of systematic effects on these scales. Note that because of the extra uncertainty introduced by this additive function, we have not used the derived E -mode signal to fit cosmological parameters. We instead directly fit the measurements

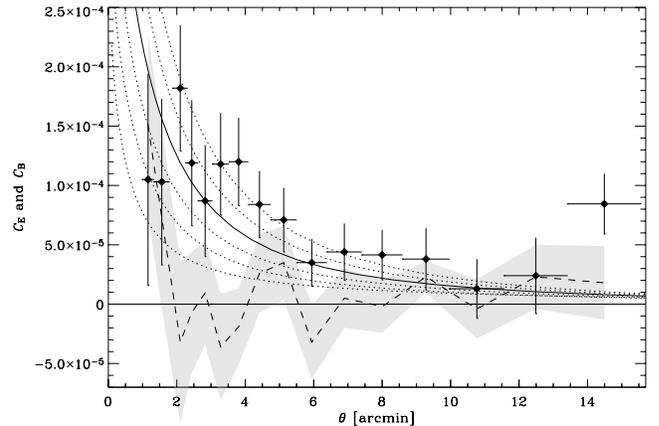


Figure 9. E - B decomposition of the shear field observed in our WHT survey. The points show the measured E (tangential) modes of the shear field. As in Fig. 3, the theoretical curves are predictions for a Λ CDM cosmology, with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $\Gamma = 0.21$ and σ_8 ranging from 0.7 (bottom) to 1.2 (top). The dashed line shows our measured B -mode (curl) signal, and the shaded region shows its 1σ error bar. In the absence of systematic effects, the B -mode should be consistent with zero.

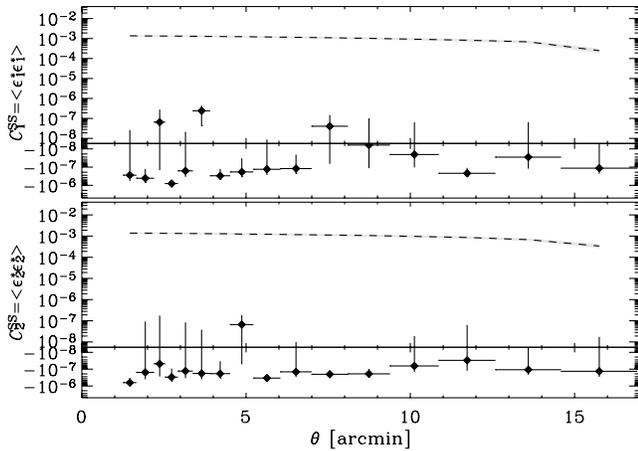


Figure 10. Correlation functions of stellar ellipticities before (dashed line) and after correction (solid points). The shaded region shows the size of 1σ errors on the pre-correction correlation functions. Application of the KSB method has successfully reduced these by approximately four orders of magnitude.

of C_1 and C_2 on those scales deemed free of systematic errors only (see Section 3.2).

4.2 PSF correction

The WHT PSF over long exposures can be quite anisotropic, with a mean stellar ellipticity of 0.051 ± 0.28 , where the error quoted is the rms stellar ellipticity within one field, averaged over all fields. Application of KSB reduces this to 0.0056 ± 0.0012 . Fig. 10 shows the full correlation functions

$$C_i^{SS} \equiv \langle e_i^* e_i^* \rangle, \quad (20)$$

where $i = \{1, 2\}$. Application of the KSB method has successfully reduced these by four orders of magnitude. Note that these are correlations of ellipticity rather than shear, and therefore exhibit a different overall normalization that is connected to the shear susceptibility factor. The elimination of stellar anisotropy is also a slightly different problem from the challenging correction of galaxy shapes, where the interpolation of the PSF is critical, and the size of the weight function r_g may be different.

The correction of galaxy shapes for the PSF anisotropy can be tested using cross-correlation functions between corrected shears γ_i from galaxies and stellar ellipticity e_i^* before correction,

$$C_i^{SG} \equiv \frac{\langle \gamma_i e_i^* \rangle^2}{\langle e_i^* e_i^* \rangle}, \quad (21)$$

where $i = \{1, 2\}$. These cross-correlation functions are shown as dashed line in the top two panels of Fig. 3, and is consistent with zero on all scales. Note that PSF correction residuals would also have appeared as B -modes in Fig. 9, which are in fact consistent with zero.

Uncertainties still remain in the KSB method concerning the overall calibration of shear estimators after PSF correction (e.g. Bernstein & Jarvis 2002; Van Waerbeke et al. 2002). Detailed image simulations by Bacon et al. (2001) or Erben et al. (2001) have been used to study these issues. Bacon et al. (2001) found that, with our survey and telescope parameters, our implementation of KSB requires a constant calibration factor of $(0.85 \pm 0.04)^{-1}$ to be applied to all calculated shear estimators.

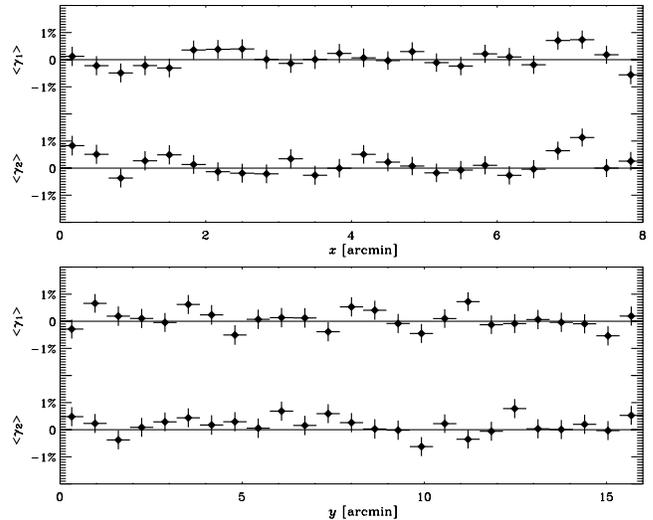


Figure 11. Galaxy shears as a function of position on the WHT CCD, averaged over all the fields used in our survey. That this is consistent with zero everywhere demonstrates an absence of systematic effects concerning CCD readout, telescope vibration and tracking.

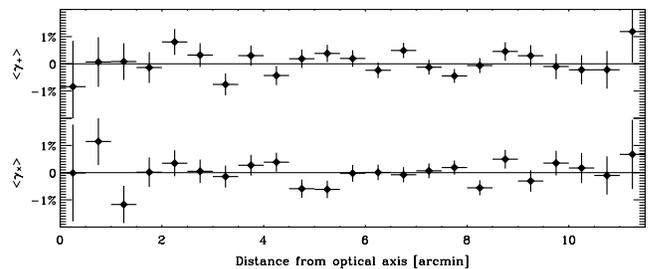


Figure 12. The radial and ‘curl’ components of galaxy shears as a function of distance from the field centre/optical axis, averaged over all the fields used in our survey. That this is consistent with zero on all scales demonstrates an absence of systematic effects concerning alt-az tracking, and the successful correction of astrometric distortions within the telescope, where the expected behaviour is purely radial (see Section 2).

4.3 Shear as a function of CCD position

Problems with read noise or charge transfer efficiency on the CCD, or telescope flexure and vibration, could cause the measured shear to vary as a function of position on the chip, even when averaged over many separate fields. Fig. 11 shows plots of shear as a function of x and y . Fig. 12 shows plots of shear as a function of r . Both are consistent with zero, as desired.

The mean components of shear across the entire CCD are $\langle \gamma_1 \rangle = (1.1 \pm 7.1) \times 10^{-4}$ and $\langle \gamma_2 \rangle = (15.6 \pm 7.0) \times 10^{-4}$. The rms shear within the survey is $\sigma_{\gamma_1} = 0.293$ and $\sigma_{\gamma_2} = 0.292$, or $\sigma_{|\gamma|} = \langle |\gamma|^2 \rangle^{1/2} = 0.413$. (Note that our shear measurement pipeline includes a catalogue cut at $|\varepsilon| < 0.5$.) The main, and irreducible, component of this dispersion comes from the intrinsic ellipticities of source galaxies. From other work performed with high-resolution and high signal-to-noise ratio space-based data (Refregier et al. 2002; Rhodes et al. 2005) and simulated images (Massey et al. 2004a), we estimate a fundamental lower limit for $\sigma_{|\gamma|}$ of around 0.30.

5 AN INDEPENDENT ANALYSIS

Our measurement of σ_8 is at the relatively high end of the distribution of published cosmic shear results. Compared with equivalently deep surveys, it is most different from that published by the COMBO-17 survey by Brown et al. (2003), who obtain $\sigma_8(\Omega_m/0.3)^{0.49} = 0.72^{+0.08}_{-0.09}$. We have therefore reanalysed our WHT data using several variations of the COMBO-17 pipeline to determine the extent to which this disagreement might arise from technical differences in the pipelines.

There are five main differences between the COMBO-17 and the WHT pipelines. In the COMBO-17 pipeline we have the following.

(i) The PSF was interpolated across every field via third-order polynomials in x and y , so the correction may be different in individual cases. Data was excluded on all scales for any fields with problematic PSF correction.

(ii) A less stringent signal-to-noise ratio cut was applied. All galaxies with $\nu > 5$ were included, raising the overall number density to 19.8. Every galaxy was still given the same weight when their shears were combined.

(iii) Galaxy pairs were not included if the galaxies lay on different CCDs. This will lower the signal-to-noise ratio on large scales relative to the WHT pipeline but may remove some bias if the chips are imperfectly aligned on the focal plane.

(iv) No correction was made for astrometric distortions in the telescope. This will spuriously increase the signal on large scales, where galaxies are a long way from the telescope axis.

(v) Independently developed code was used to calculate theoretical models and to fit cosmological parameters to the data.

5.1 Shear measurement

The first two changes are relevant in the creation of a shear catalogue. Differences in the PSF interpolation introduce a ~ 1 per cent rms dispersion between shear values measured by the WHT or COMBO-17 pipelines (see the top-middle panel in Fig. 13). This small difference can be explained by the uncertain nature of any interpolation between sparsely sampled data points. The lower-order fit in the WHT method typically produces smoother, and therefore perhaps more reasonable behaviour in regions of the images that

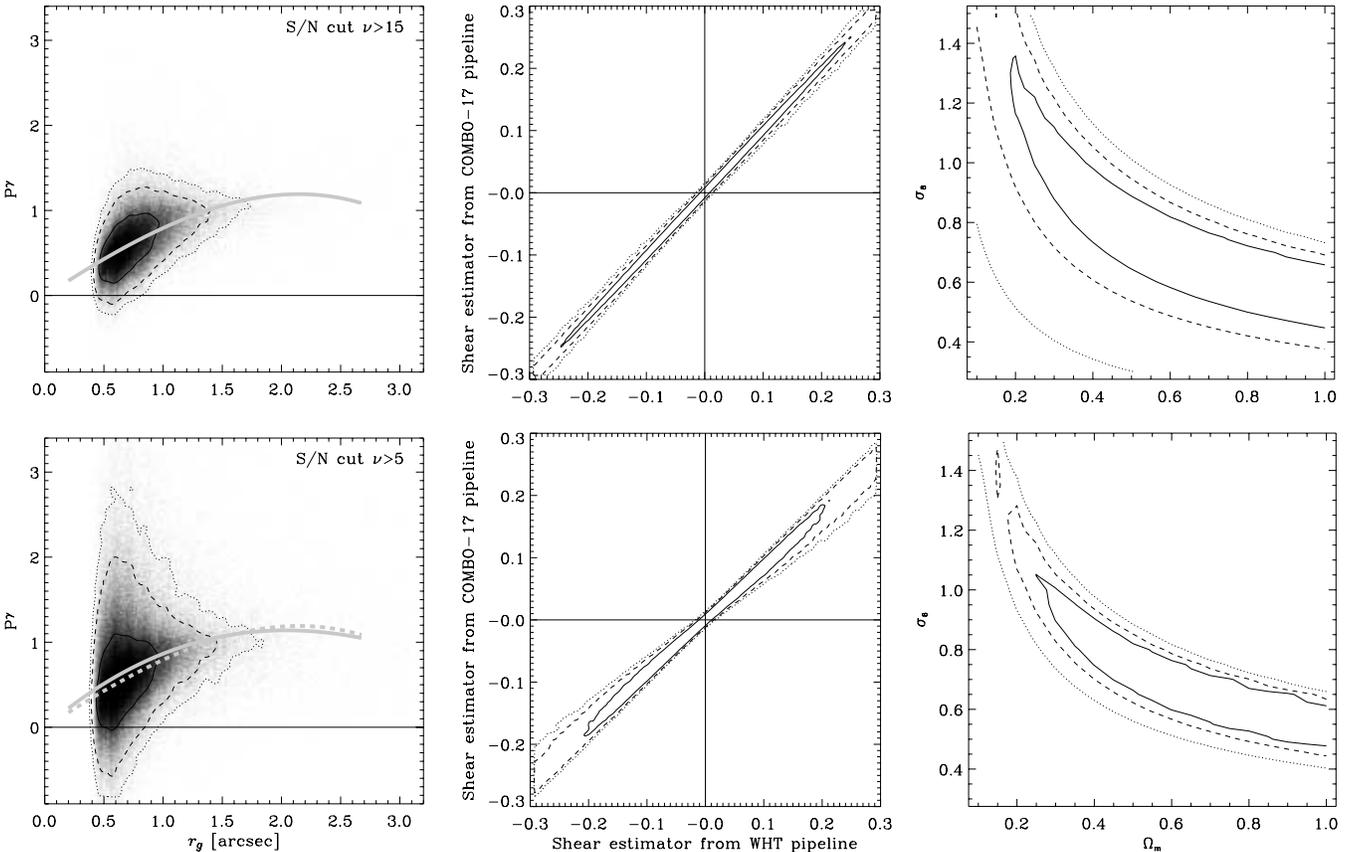


Figure 13. Left-hand panels: the shear susceptibility factor P^γ for all galaxies, as a function of their size r_g , with different options for the signal-to-noise ratio cut in the catalogue. The grey-scale shows the number density of galaxies throughout the parameter space, calculated by the COMBO-17 pipeline. The WHT pipeline produces results almost identical to the top-left panel. In practice the fit is performed on galaxies for one image at a time, and is therefore more noisy; the fit to all galaxies here is shown merely to guide the eye. The fit in the top panel is reproduced as a dashed line in the bottom panel. When the COMBO-17 pipeline uses a signal-to-noise ratio cut of $\nu > 5$, the average shear susceptibility of the population ensemble is raised. Middle panels: a comparison of the shear estimators derived via the WHT and the COMBO-17 shear measurement pipelines. Only subsequently matched galaxies are included in this plot, and the WHT pipeline always uses a signal-to-noise ratio cut to the catalogue of $\nu > 15$, so all galaxies shown are brighter than that. Consistent results are obtained in the top panel, where the COMBO-17 pipeline also uses a cut of $\nu > 15$. The unbiased, 1 per cent dispersion of individual values is merely due to the different interpolation of the PSF across individual fields. When the COMBO-17 pipeline uses a signal-to-noise ratio cut of $\nu > 5$, the shear estimators of all galaxies are lowered, including these bright ones. Right-hand panels: constraints upon cosmological parameters Ω_m and σ_8 , from a maximum-likelihood analysis of shears measured with the COMBO-17 pipeline, using different options for the signal-to-noise ratio cut. The 68.3, 95.4 and 99.7 per cent confidence contours are shown.

are devoid of stars or near the edge of the CCD. However, neither method appears to bias the shear estimators or resulting constraints on cosmological parameters.

Of greater importance are the $\nu > 15$ or $\nu > 5$ catalogue cuts in signal-to-noise ratio. Very faint galaxies were discarded in our original analysis of WHT images, and the simulated images of Bacon et al. (2001), because of residual correlation found between the galaxy and stellar ellipticities after correction. Such correlations were not observed in the COMBO-17 data, so the cut was lowered. A possible explanation is the absence of CCD fringing in the COMBO-17 data. Uncorrected fringing in WHT images would alter the shapes of those galaxies at a flux level comparable to the fringing via an additive process rather than a convolution or multiplication. This would not have been corrected, so the galaxies would have simply been discarded. Brighter galaxies that appear in both the WHT-pipeline and COMBO-17 pipeline have identical estimates of P^ν , since both are calculated by HFINDPEAKS. However, these values are noisy. Both pipelines regard galaxies as a population ensemble, for which it is possible to average over noise by fitting a global shear susceptibility factor.

Theoretically, the susceptibility factor of a galaxy shear is expected to vary as a function of its radial profile, size and ellipticity; but not as a function of its flux. We therefore fit $P^\nu = P^\nu(r_g)$, as shown in the left-hand panels of Fig. 13. However, it seems that real faint galaxies do indeed have higher shear susceptibilities than similarly sized bright galaxies. As shown in the middle panels of Fig. 13, this raises $P^\nu(r_g)$, and thus lowers shear estimators by a factor of $0.85^{+0.12}_{-0.5}$. The skewed distribution reflects the overall size distribution of galaxies, and arises because this process affects small galaxies more than large ones. The reason for this variation of $P^\nu(\nu)$ is not yet clear: it may be that the morphology distribution of faint galaxies really does contain intrinsically higher shear susceptibilities. Both approaches would be valid in this case, since the fit really has found a suitable average value for the population ensemble of galaxies.

A more worrying alternative would be that the noise in faint galaxies biases P^ν to higher values, partly because the main P^{sh} component of P^ν (see Kaiser et al. 1995) is defined to be strictly positive. In this case, it would instead be preferable to fit only the bright galaxies, and then apply that susceptibility to faint galaxies of the same size. Either way, the results from the WHT pipeline would be less affected, as the $\nu > 15$ cut has been calibrated upon simulated images containing a known signal by Bacon et al. (2001).

A full investigation into this technical issue is beyond the scope of this paper, and it will need further investigation in the future. Other comparative studies, e.g. the Shear TEsting Program (STEP, Heymans et al. in preparation), or the Edinburgh and Bonn pipelines (Heymans et al. in preparation), find that the detailed way in which P^ν is, or is not, adequately fitted may indeed explain a large part of the variation between results from different implementations of KSB. Here, we shall propagate the analysis of shear catalogues from the COMBO-17 pipeline with both $\nu > 15$ and $\nu > 5$ cuts.

5.2 Cosmological parameter constraints

We find similar instabilities in the production of shear–shear correlation functions with the COMBO-17 pipeline as we found in the WHT pipeline; results are sensitive to binning at the $\sim \pm 7$ per cent level. Measurements with $\nu > 5$ are shown as blue circles in Fig. 3, rescaled by a factor of $(0.8/0.9)^{1.65}$ to allow a comparison despite the increased source redshift distribution. Excluding galaxy pairs of different CCDs has increased the size of error bars by up to

50 per cent at large θ , where the number of available galaxy pairs is significantly lower, but the central values move only within their 1σ error bars. Not correcting for the known astrometric distortion of the telescope has moved the bin at largest scales by a significant amount, but this is not used for parameter fitting anyway. Excluding data on all scales from fields with any residual star–galaxy correlation also has a minimal effect. Each of these differences change the derived constraint on σ_8 by less than 1 per cent.

The main difference in the two analyses is that introduced by the different cuts in signal-to-noise ratio. The inclusion of fainter galaxies increases the median magnitude of the $\nu > 5$ population by 0.84 ± 0.4 mag, leading to a new median redshift of $z_s = 0.90 \pm 0.07$ (Cohen et al. 2000). The expected cosmic shear signal thus increases (see equation 14) – an effect that is incorporated during the calculation of theoretical models for the fitting of cosmological parameters. Reassuringly, the shear signal measured from faint galaxies is also higher. The final constraints for the COMBO-17 pipeline, using a cut of $\nu > 15$ indicate:

$$\sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{0.52} = 1.01 \pm 0.13, \quad (22)$$

including only statistical errors and, for a cut of $\nu > 5$ with the higher assumed source redshift distribution

$$\sigma_8 \left(\frac{\Omega_m}{0.3} \right)^{0.52} = 0.98 \pm 0.10. \quad (23)$$

The consistency of these fits, and their agreement with the results from the WHT pipeline is reassuring. They tend to justify the adoption of a global shear susceptibility factor for an ensemble population of galaxies, and verify the validity of the methods for controlling many potential systematic effects in both pipelines.

Our survey probes a wide range of angular scales, and thus excludes small values of $\Omega_m < 0.25$ (at 68 per cent CL). The preference for large Ω_m comes from the fit to the transfer function by Bardeen et al. (1986). If we adopt the fit by Eisenstein & Hu (1997), and assume the Hubble parameter $h = 0.7$, we effectively apply a different prior on Ω_m . In this case, our data excludes $\Omega_m > 0.55$ (at 68 per cent CL), similar to the result obtained by Van Waerbeke et al. (2005). Using the Eisenstein & Hu (1997) transfer function also lowers our best-fitting value for σ_8 around $\Omega_m = 0.3$ by ~ 2 per cent.

6 CONCLUSIONS

We have measured the weak-lensing shear–shear correlation functions in 4 deg^2 of deep R -band imaging data from the William Herschel Telescope. Our measurements constrain the amplitude of the mass power spectrum, $\sigma_8(\Omega_m/0.3)^{0.52} = 1.02 \pm 0.15$, including all contributions to the total 68 per cent CL error budget: statistical noise, instabilities due to non-Gaussian noise, sample variance, covariance between different angular scales, systematic measurement and detection biases, source redshift uncertainty and marginalization with priors over other parameters. We have examined our data for contamination by systematic effects using a variety of tests including an E – B decomposition. These demonstrate a well-understood and modest contribution to our uncertainties from systematic errors.

Using the pipeline developed for earlier WHT studies, we find our measurement of the normalization of the dark matter power spectrum lies at the relatively high end of the distribution of published values. However, it is still consistent with those from equivalently deep surveys by Refregier et al. (2002) and Rhodes et al. (2005). Our results are also consistent at the 1σ level with

cosmic microwave background (CMB) results from the *Wilkinson Microwave Anisotropy Probe* (WMAP; Spergel et al. 2003).

The wide distribution of σ_8 constraints from recent cosmic shear surveys understandably casts some aspersions upon their precision. For example, it might be argued that the dispersion largely arises from unknown or poorly understood systematic effects. For the first time we analyse our data set with an independent pipeline – that developed for the COMBO-17 data (Brown et al. 2003). This provides a valuable check on the extent to which dispersion in the published cosmological results, and our apparently high normalization, might arise from different techniques, both in constructing shear catalogues and in analysing the shear–shear correlation functions. Reassuringly, we find remarkable concordance between the two pipelines for the same data set. However, the various differences we have explored, for example in the selection thresholds, seem to be insufficient to reconcile the spread in observed σ_8 values, suggesting that significant differences remain at some level in the actual data themselves.

Uncertainties in the redshift distribution of source galaxies in deep data clearly contribute. It is difficult to determine the precise redshift distribution of galaxies after excluding those smaller than a fixed apparent size. We have been conservative in this analysis and, as seen in equation (15), source redshift uncertainty is already a major component of our total error budget. The resolution of such issues will require extensive spectroscopic follow-up and more complete image simulations. Such advances are essential if the potential of the next generation of cosmic shear surveys is to be fully realized.

Finally, we note that most recent cosmic shear results remain discrepant at the 3σ level with measurements derived from the abundance of X-ray-selected cluster samples based on an observational rather than theoretical mass–temperature relation (Borgani et al. 2001; Seljak 2001; Reiprich & Böhringer 2002; Viana et al. 2002). These suggest $\sigma_8 \approx 0.75$. Amara & Refregier (2004) concluded that even extreme non-Gaussianity in the mass distribution would be insufficient to explain this discrepancy, because the two techniques probe similar mass scales. Further studies are therefore needed in both the cluster method, to understand the difference between the observed mass–temperature relation and that found in numerical simulations; and in the weak-lensing method, to construct more reliable and better calibrated shear measurement methods. Such consistency checks will represent a crucial verification of the standard Λ CDM paradigm, so resolving this issue is of paramount importance.

ACKNOWLEDGMENTS

We thank Neil O’Mahony, Chris Benn and the WHT staff for their help with the observations. We thank Nick Kaiser for providing us with the IMCAT software, and to Douglas Clowe for advice on its use. We thank Sarah Bridle, Catherine Heymans and Jason Rhodes for useful discussions. We thank the anonymous referee for helpful suggestions to improve the text and to clarify various discussions. DB and MB were supported by PPARC post-doctoral fellowships.

REFERENCES

Amara A., Refregier A., 2004, *MNRAS*, 351, 375
 Bacon D., Refregier A., Ellis R., 2000, *MNRAS*, 318, 625

Bacon D., Refregier A., Clowe D., Ellis R., 2001, *MNRAS*, 325, 1065
 Bacon D., Massey R., Refregier A., Ellis R., 2003, *MNRAS*, 344, 673
 Bardeen J., Bond J., Kaiser N., Szalay A., 1986, *ApJ*, 304, 15
 Bartelmann M., Schneider P., 2000, *Phys. Rep.*, 340, 291
 Bernardeau F., 1999, *NATO Sci. Ser. C*, 541, 179
 Bernstein G., Jarvis M., 2002, *ApJ*, 123, 583
 Borgani S. et al., 2001, *ApJ*, 561, 13
 Brown M., Taylor A., Bacon D., Gray M., Dye S., Meisenheimer K., Wolf C., 2003, *MNRAS*, 341, 100
 Cohen J., Hogg D., Blandford R., Hu E., Songaila A., Shopbell P., Richberg K., 2000, *ApJ*, 538, 29
 Crittenden R., Natarajan P., Pen U.-L., Theuns T., 2000, *ApJ*, 559, 552
 Dekel A., Lahav O., 1999, *ApJ*, 520, 24
 Eisenstein D., Hu W., 1997, *ApJ*, 511, 5
 Erben T., van Waerbeke L., Bertin E., Mellier Y., Schneider P., 2001, *A&A*, 366, 717
 Gray M., Taylor A., Meisenheimer K., Dye S., Wolf C., Thommes E., 2002, *ApJ*, 568, 141
 Hamana T. et al., 2003, *ApJ*, 597, 98
 Hämmerle et al., 2002, *A&A*, 385, 743
 Hoekstra H., 2004, *MNRAS*, 347, 1337
 Hoekstra H., Yee H., Gladders M., Felipe Barrientos L., Hall P., Infante L., 2002a, *ApJ*, 572, 55
 Hoekstra H., van Waerbeke L., Gladders M., Mellier Y., Yee H., 2002b, *ApJ*, 577, 604
 Huterer D., White M., 2003, *ApJ*, 578, L95
 Jarvis M., Jain B., 2005, *ApJ*, submitted (astro-ph/0412234)
 Jarvis M. et al., 2003, *AJ*, 125, 1014
 Kaiser N., Squires G., Broadhurst T., 1995, *ApJ*, 449, 460 (KSB)
 Massey R., Bacon D., Refregier A., Ellis R., 2001, in Shanks T., Metcalfe N. eds, *ASP Conf. Ser. Vol. 283. Astron. Soc. Pac.*, San Francisco, p. 193
 Massey R. et al., 2004a, *AJ*, 127, 3089
 Massey R., Refregier A., Conselice C., Bacon D., 2004b, *MNRAS*, 348, 214
 Mellier Y., 1999, *ARA&A*, 37, 127
 Peacock J., Dodds S., 1996, *MNRAS*, 280, 19
 Pen U., Van Waerbeke L., Mellier Y., 2002, *ApJ*, 567, 31
 Percival W. et al., 2001, *MNRAS*, 327, 1297
 Pierpaoli E., Scott D., White M., 2001, *MNRAS*, 325, 77
 Refregier A., 2003, *ARA&A*, 41, 645
 Refregier A., Rhodes J., Groth E., 2002, *ApJ*, 572, 131
 Reiprich T., Böhringer H., 2002, *ApJ*, 567, 716
 Rhodes J., Refregier A., Groth E., 2001, *ApJ*, 552, 85
 Rhodes J., Refregier A., Collins N., Gardner J., Groth E., Hill R., 2005, *ApJ*, 605, 29
 Schneider P., van Waerbeke L., Mellier Y., 2002, *A&A*, 389, 729
 Seljak U., 2001, *MNRAS*, 337, 769
 Smith G., Edge A., Eke V., Nichol R., Smail I., Kneib J.-P., 2003a, *AJ*, 590, L79
 Smith R. et al., 2003b, *MNRAS*, 341, 1311
 Spergel D. et al., 2003, *ApJS*, 148, 175
 Szalay A. et al., 2003, *ApJ*, 591, 1
 van Waerbeke L., Mellier Y., Pelló R., Pen U.-L., McCracken H., Jain B., 2002, *A&A*, 393, 369
 van Waerbeke L., Mellier Y., Hoekstra H., 2005, *A&A*, 429, 75
 Viana P., Nichol R., Liddle A., 2002, *ApJ*, 569, 75
 Weinberg D., Davé R., Katz N., Hernquist L., 2004, *ApJ*, 601, 1
 Wittman D., 2002, *Lecture Notes in Physics*, 608, 55

This paper has been typeset from a $\text{\TeX}/\text{\LaTeX}$ file prepared by the author.